



Financial Econometrics and Empirical Finance II

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1 Dynamic Models

```
"""
@author: dtoppo
Dynamic models
"""

import numpy as np
import matplotlib.pyplot as plt
import pandas as pd
from statsmodels.formula.api import ols
from statsmodels.stats.anova import anova_lm

# Import dataserie from Excel
dataFile = "data_python_HW_2017.xlsx"
xlsx = pd.ExcelFile(dataFile)

sheetName = "Dynamic Models"
dataSeries = xlsx.parse(sheetName)

# =====
# Regression result
# dir(model) to display model attributes
# =====
model = ols("USVW ~ Treasury", dataSeries).fit()
print(model.summary()) # table summary
```

OLS Regression Results

Dep. Variable:	y	R-squared:	0.046
Model:	OLS	Adj. R-squared:	0.040
Method:	Least Squares	F-statistic:	7.409
Date:	Thu, 25 May 2017	Prob (F-statistic):	7.42e-05
Time:	11:01:50	Log-Likelihood:	755.82
No. Observations:	465	AIC:	-1504.
Df Residuals:	461	BIC:	-1487.
Df Model:	3		
Covariance Type:	nonrobust		

	coef	std err	t	P> t	[0.025	0.975]
Intercept	0.0093	0.002	3.952	0.000	0.005	0.014
x1	0.2155	0.047	4.630	0.000	0.124	0.307
x2	-0.0317	0.048	-0.667	0.505	-0.125	0.062
x3	-0.0318	0.046	-0.688	0.492	-0.123	0.059

Omnibus:	49.632	Durbin-Watson:	1.986
Prob(Omnibus):	0.000	Jarque-Bera (JB):	238.733
Skew:	-0.287	Prob(JB):	1.44e-52
Kurtosis:	6.463	Cond. No.	24.4

Figure 1: OLS Regression Results

```

#
# Autocorrelation analysis
#
residuals = model.resid
lagged_residuals = model.resid.shift()
corr_matrix = np.corrcoef(residuals[1:], lagged_residuals[1:])
print(corr_matrix)

```

$$\begin{bmatrix} 1 & 0.075 \\ 0.075 & 1 \end{bmatrix}$$

```

#
# Plot residuals
#
x_range = np.arange(0, np.size(residuals))
plt.plot(x_range, residuals, 'bD')

hlines_range = np.arange(-0.3, 0.2, 0.05)
for hline in hlines_range:
    plt.axhline(y=hline, color='#aaaaaa', linestyle=":")

min_value = -0.3
max_value = 0.15

plt.ylim(min_value, max_value)

```

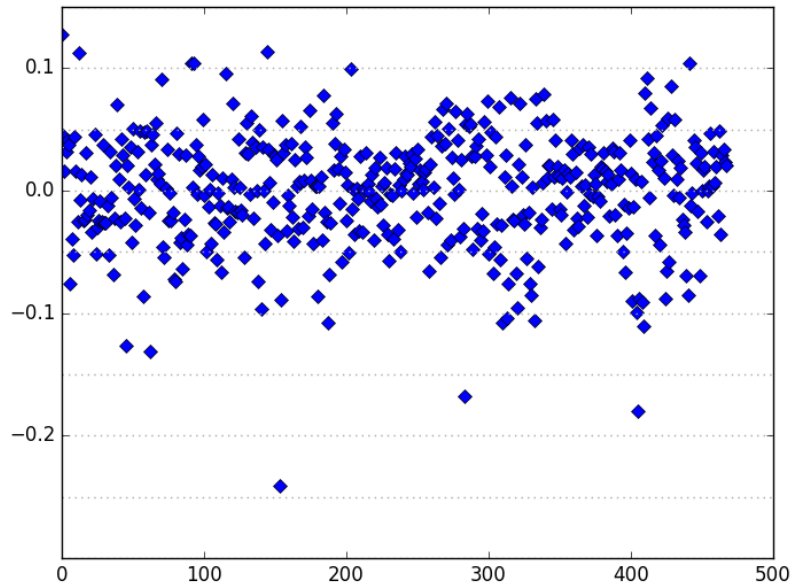


Figure 2: Residuals

Looking at the values of the residuals and at the graph above, there seems to be a tendency for negative values to follow negative values, and positive values to follow positive values. This is consistent with positive correlation between successive terms.

We can also check numerically the correlation between the residuals and the lagged values. The output indicates about 0.075 correlation between the errors one period apart.

2 ARMA Models

```
"""
@author: dtoppo
ARMA models
"""

import numpy as np
import matplotlib.pyplot as plt
import pandas as pd
from statsmodels.formula.api import ols
from scipy import optimize

# Import dataserie from Excel
dataFile = "data_python-HW-2017.xlsx"
xlsx = pd.ExcelFile(dataFile)

sheetName = "ARMA"
dataSeries = xlsx.parse(sheetName)

ret_lag_1 = dataSeries.shift()[3:]
ret_lag_2 = dataSeries.shift(2)[3:]
ret_lag_3 = dataSeries.shift(3)[3:]

#-----
# Multiple regression
# dir(model) to display model attributes
#-----
serie = "Belgium"
y = dataSeries[serie][3:]
data = pd.concat([y, ret_lag_1[serie], ret_lag_2[serie], ret_lag_3[serie]],
                 axis=1, keys=["y", "x1", "x2", "x3"])
model = ols("y ~ x1 + x2 + x3", data).fit()
print(model.summary())
```

OLS Regression Results						
Dep. Variable:	y	R-squared:	0.046			
Model:	OLS	Adj. R-squared:	0.040			
Method:	Least Squares	F-statistic:	7.409			
Date:	Thu, 25 May 2017	Prob (F-statistic):	7.42e-05			
Time:	14:31:20	Log-Likelihood:	755.82			
No. Observations:	465	AIC:	-1504.			
Df Residuals:	461	BIC:	-1487.			
Df Model:	3					
Covariance Type:	nonrobust					
	coef	std err	t	P> t	[0.025	0.975]
Intercept	0.0093	0.002	3.952	0.000	0.005	0.014
x1	0.2155	0.047	4.630	0.000	0.124	0.307
x2	-0.0317	0.048	-0.667	0.505	-0.125	0.062
x3	-0.0318	0.046	-0.688	0.492	-0.123	0.059
Omnibus:	49.632	Durbin-Watson:	1.986			
Prob(Omnibus):	0.000	Jarque-Bera (JB):	238.733			
Skew:	-0.287	Prob(JB):	1.44e-52			
Kurtosis:	6.463	Cond. No.	24.4			

Figure 3: OLS Regression Results

```

#
# Plots
#
n = np.size(y)
x_range = np.arange(0, n)

predictedY = model.predict() # predicted values
residuals = model.wresid # residuals
returns = predictedY + residuals

# Plot data
plt.plot(x_range, predictedY)
plt.plot(x_range, returns)

```

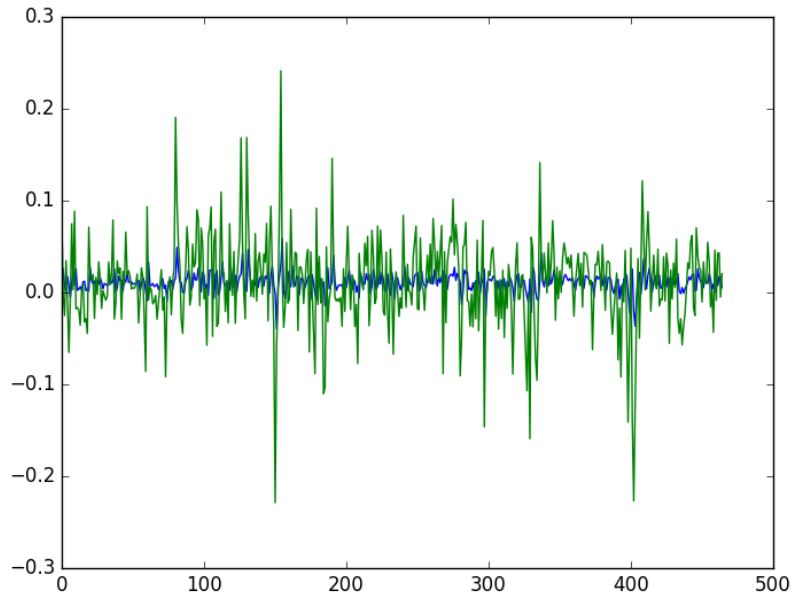


Figure 4: Actual and Fitted Values

```

# Residual plots
plt.figure()
plt.subplot(311)
plt.plot(ret_lag_1[serie], residuals, "p")
plt.axhline(y=0, color='#aaaaaa', linestyle=":")
plt.axvline(x=0, color='#aaaaaa', linestyle=":")

plt.subplot(312)
plt.plot(ret_lag_2[serie], residuals, "p")
plt.axhline(y=0, color='#aaaaaa', linestyle=":")
plt.axvline(x=0, color='#aaaaaa', linestyle=":")

plt.subplot(313)
plt.plot(ret_lag_3[serie], residuals, "p")
plt.axhline(y=0, color='#aaaaaa', linestyle=":")
plt.axvline(x=0, color='#aaaaaa', linestyle=":")

```

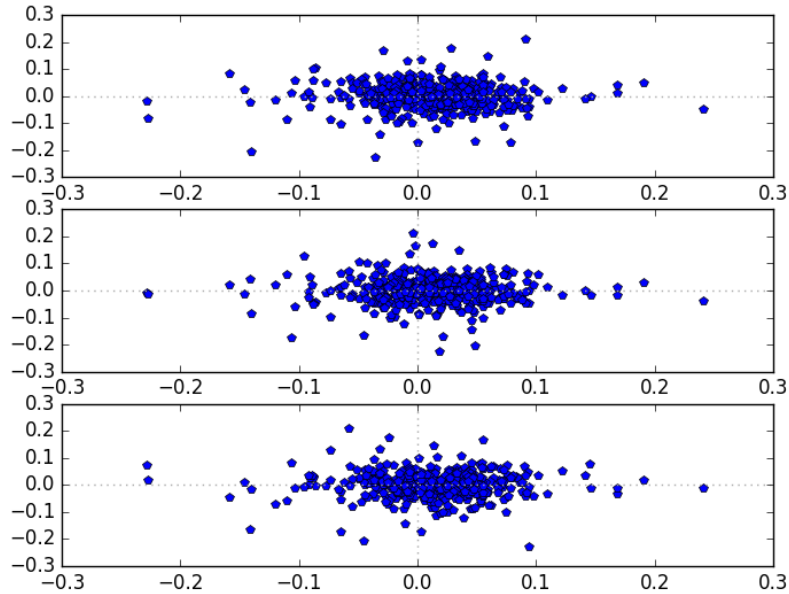


Figure 5: Residual Plots

```

# Fit plots
plt.figure()
plt.subplot(311)
plt.plot(ret_lag_1[serie], predictedY, "p")
plt.plot(ret_lag_1[serie], y, "p")
plt.axhline(y=0, color='#aaaaaa', linestyle=":")
plt.axvline(x=0, color='#aaaaaa', linestyle=":")

plt.subplot(312)
plt.plot(ret_lag_2[serie], predictedY, "p")
plt.plot(ret_lag_2[serie], y, "p")
plt.axhline(y=0, color='#aaaaaa', linestyle=":")
plt.axvline(x=0, color='#aaaaaa', linestyle=":")

plt.subplot(313)
plt.plot(ret_lag_3[serie], predictedY, "p")
plt.plot(ret_lag_3[serie], y, "p")
plt.axhline(y=0, color='#aaaaaa', linestyle=":")
plt.axvline(x=0, color='#aaaaaa', linestyle=":")

```

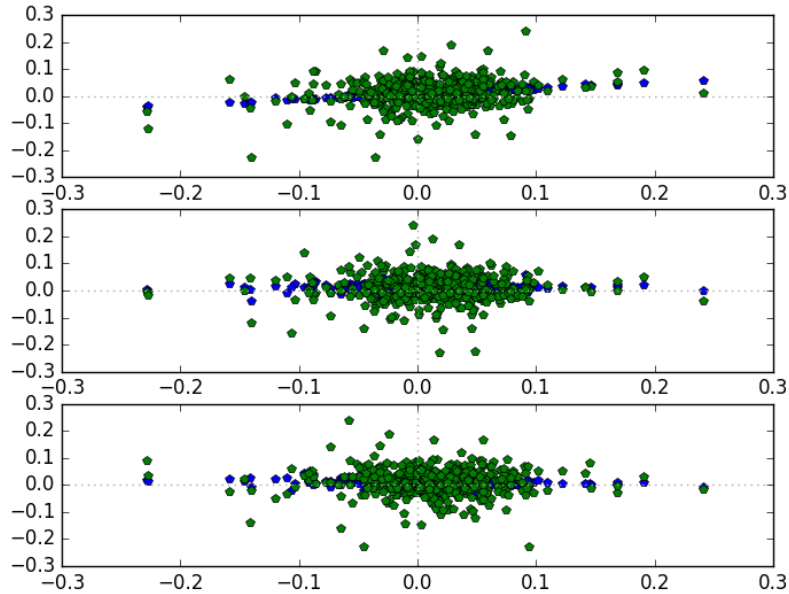


Figure 6: Line Fit Plots

```

#
# ARIMA(1, 1, 0)
#
serie = "France"
returns = dataSeries[serie]

# ARIMA(1, 1, 0)
def arimal(params):
    rho, theta = params

    n = np.size(returns)
    ma = np.zeros(n)
    ma[0] = 0

    for i in range(1, n):
        ma[i] = returns[i] - (rho * returns[i-1] + theta * ma[i-1])

    return ma

# Squared errors function to be minimized
def squaredError(params):
    return (arimal(params)**2).sum()

initParams = [0, 0]

results = optimize.minimize(squaredError, initParams, method='SLSQP')
print(results.x)

```


3 Simultaneous Equations

```
"""
@author: dtoppo
Simultaneous Equations
"""

import scipy as sp
import pandas as pd
from statsmodels.formula.api import ols

# Import dataserie from Excel
dataFile = "data_python_HW_2017.xlsx"
xlsx = pd.ExcelFile(dataFile)

sheetName = "Simult Equ"
dataSeries = xlsx.parse(sheetName)

# =====
# Descriptive statistics
# =====

def descriptiveStats(serieName):

    serie = dataSeries[serieName]

    mean = sp.mean(serie)
    stdErr = sp.stats.sem(serie)
    median = sp.median(serie)
    mode = sp.stats.mode(serie)
    stdDev = sp.std(serie, ddof=1)
    var = sp.var(serie, ddof=1)
    kurtosis = sp.stats.kurtosis(serie, fisher=True)
    skewness = sp.stats.skew(serie)
    minVal = sp.amin(serie)
    maxVal = sp.amax(serie)
    rangeVal = maxVal - minVal
    sumVal = sp.sum(serie)
    count = sp.count_nonzero(serie)
    confInterval = sp.stats.norm.interval(0.05, loc=mean, scale=stdDev)

    print("*****")
    print("—— %s ——" %serieName)
    print("Mean: %.2f" %mean)
    print("Standard Error: %.2f" %stdErr)
    print("Median: %.2f" %median)
    print("Mode: %.2f" %mode[0])
    print("Standard Deviation: %.2f" %stdDev)
    print("Sample Variance: %.2f" %var)
    print("Kurtosis: %.2f" %kurtosis)
    print("Skewness: %.2f" %skewness)
    print("Range: %.2f" %rangeVal)
    print("Minimum: %.2f" %minVal)
    print("Maximum: %.2f" %maxVal)
    print("Sum: %.2f" %sumVal)
    print("Count: %.2f" %count)
    print("Confidence Level (95%%): %.2f" %(confInterval[1] - confInterval[0]))
    print("*****")
```

```

for serieName in dataSeries:
    descriptiveStats(serieName)

*****

--- Returns ---
Mean: 0.01
Standard Error: 0.00
Median: 0.02
Mode: 0.02
Standard Deviation: 0.04
Sample Variance: 0.00
Kurtosis: 3.40
Skewness: -1.00
Range: 0.36
Minimum: -0.23
Maximum: 0.13
Sum: 2.52
Count: 253.00
Confidence Level (95%): 0.01
*****

```

Figure 7: Summary Statistics for Returns

```

#-----#
# Regressions
#-----#
data = pd.concat([dataSeries["Returns"], dataSeries["dprod"], dataSeries["dspread"],
                 dataSeries["rterm"], dataSeries["dcredit"], dataSeries["dmoney"]], axis=1)
returns_model = ols("Returns ~ dprod + dspread + rterm + dcredit + dmoney", data).fit()
print(returns_model.summary())

```

OLS Regression Results						
Dep. Variable:	Returns	R-squared:	0.010			
Model:	OLS	Adj. R-squared:	-0.011			
Method:	Least Squares	F-statistic:	0.4745			
Date:	Thu, 25 May 2017	Prob (F-statistic):	0.795			
Time:	14:52:11	Log-Likelihood:	433.79			
No. Observations:	253	AIC:	-855.6			
Df Residuals:	247	BIC:	-834.4			
Df Model:	5					
Covariance Type:	nonrobust					
	coef	std err	t	P> t	[0.025	0.975]
Intercept	0.0113	0.005	2.474	0.014	0.002	0.020
dprod	-0.0015	0.002	-0.994	0.321	-0.005	0.002
dspread	-0.0396	0.042	-0.943	0.346	-0.122	0.043
rterm	0.0034	0.010	0.320	0.749	-0.017	0.024
dcredit	-1.706e-07	5.06e-07	-0.337	0.736	-1.17e-06	8.25e-07
dmoney	-0.0012	0.002	-0.546	0.586	-0.006	0.003
Omnibus:	57.083	Durbin-Watson:	1.952			
Prob(Omnibus):	0.000	Jarque-Bera (JB):	151.874			
Skew:	-1.006	Prob(JB):	1.05e-33			
Kurtosis:	6.219	Cond. No.	1.37e+05			

Figure 8: Reduced-form equation for Returns

```
data = pd.concat([dataSeries["Inflation"], dataSeries["dprod"], dataSeries["dspread"],
                 dataSeries["rterm"], dataSeries["dcredit"], dataSeries["dmoney"]], axis=1)
inflation_model = ols("Inflation ~ dprod + dspread + rterm + dcredit + dmoney", data).fit()
print(inflation_model.summary())
```

OLS Regression Results						
Dep. Variable:	Inflation		R-squared:	0.219		
Model:	OLS		Adj. R-squared:	0.203		
Method:	Least Squares		F-statistic:	13.88		
Date:	Thu, 25 May 2017		Prob (F-statistic):	5.93e-12		
Time:	14:53:31		Log-Likelihood:	-841.69		
No. Observations:	253		AIC:	1695.		
Df Residuals:	247		BIC:	1717.		
Df Model:	5					
Covariance Type:	nonrobust					
	coef	std err	t	P> t	[0.025	0.975]
Intercept	214.3211	0.709	302.171	0.000	212.924	215.718
dprod	0.0109	0.241	0.045	0.964	-0.463	0.485
dspread	7.1308	6.495	1.098	0.273	-5.662	19.924
rterm	-1.0572	1.624	-0.651	0.516	-4.255	2.141
dcredit	0.0006	7.82e-05	8.113	0.000	0.000	0.001
dmoney	0.1552	0.346	0.448	0.654	-0.527	0.837
Omnibus:	1.397	Durbin-Watson:	0.242			
Prob(Omnibus):	0.497	Jarque-Bera (JB):	1.281			
Skew:	-0.025	Prob(JB):	0.527			
Kurtosis:	2.655	Cond. No.	1.37e+05			

Figure 9: Reduced-form equation for Inflation

The importance of using 2SLS:

If we estimate a SSSE by OLS, OLS estimators will suffer from the so-called Simultaneous Equation Bias, arising from the presence of correlation between independent variables and regressors in some equations of the SSSE. The errors will be correlated with the regressors and this violates an assumption of the regression framework. Applying standard ordinary least squares (OLS) under these circumstances results in inconsistent estimates.

To remedy this problem we can apply 2SLS: 2SLS imply replacing endogenous variables on the RHS with fitted OLS values. Even if 2SLS estimators are still biased, the advantage of 2SLS estimators over OLS is that they are consistent.

```

#-----
# 2SLS Model with predicted returns and predicted inflation
#-----
predicted_returns = pd.DataFrame({"PredictedReturns" : returns_model.predict()},
                                index=dataSeries.index)
preidcted_inflation = pd.DataFrame({"PredictedInflation" : inflation_model.predict()},
                                   index=dataSeries.index)

data = pd.concat([dataSeries["Inflation"], predicted_returns, dataSeries["dcredit"],
                 dataSeries["dprod"], dataSeries["dmoney"]], axis=1)
inflation_model = ols("Inflation ~ PredictedReturns + dcredit + dprod + dmoney", data).fit()
print(inflation_model.summary())

```

OLS Regression Results						
Dep. Variable:	Inflation	R-squared:	0.219			
Model:	OLS	Adj. R-squared:	0.206			
Method:	Least Squares	F-statistic:	17.39			
Date:	Thu, 25 May 2017	Prob (F-statistic):	1.36e-12			
Time:	14:54:29	Log-Likelihood:	-841.73			
No. Observations:	253	AIC:	1693.			
Df Residuals:	248	BIC:	1711.			
Df Model:	4					
Covariance Type:	nonrobust					
	coef	std err	t	P> t	[0.025	0.975]
Intercept	216.5030	1.946	111.274	0.000	212.671	220.335
PredictedReturns	-192.9712	156.389	-1.234	0.218	-500.991	115.048
dcredit	0.0006	8.36e-05	7.209	0.000	0.000	0.001
dprod	-0.2824	0.333	-0.847	0.398	-0.939	0.374
dmoney	-0.0948	0.372	-0.255	0.799	-0.827	0.638
Omnibus:	1.328	Durbin-Watson:	0.242			
Prob(Omnibus):	0.515	Jarque-Bera (JB):	1.233			
Skew:	-0.018	Prob(JB):	0.540			
Kurtosis:	2.660	Cond. No.	3.30e+06			

Figure 10: Inflation

```
data = pd.concat([dataSeries["Returns"], predicted_inflation, dataSeries["dprod"],
                 dataSeries["dspread"], dataSeries["rterm"]], axis=1)
inflation_model = ols("Returns ~ PredictedInflation + dprod + dspread + rterm", data).fit()
print(inflation_model.summary())
```

OLS Regression Results						
Dep. Variable:	Returns	R-squared:	0.008			
Model:	OLS	Adj. R-squared:	-0.008			
Method:	Least Squares	F-statistic:	0.5258			
Date:	Thu, 25 May 2017	Prob (F-statistic):	0.717			
Time:	15:25:47	Log-Likelihood:	433.65			
No. Observations:	253	AIC:	-857.3			
Df Residuals:	248	BIC:	-839.6			
Df Model:	4					
Covariance Type:	nonrobust					
	coef	std err	t	P> t	[0.025	0.975]
Intercept	0.0787	0.173	0.454	0.650	-0.262	0.420
PredictedInflation	-0.0003	0.001	-0.395	0.693	-0.002	0.001
dprod	-0.0016	0.002	-1.069	0.286	-0.005	0.001
dspread	-0.0360	0.042	-0.850	0.396	-0.119	0.047
rterm	0.0020	0.010	0.193	0.847	-0.018	0.022
Omnibus:	57.420	Durbin-Watson:	1.954			
Prob(Omnibus):	0.000	Jarque-Bera (JB):	157.719			
Skew:	-0.999	Prob(JB):	5.65e-35			
Kurtosis:	6.312	Cond. No.	1.37e+04			

Figure 11: Returns

Except for dcredit in the Inflation equation and intercept in Returns Equation, none of the parameters is even close to statistical significance. The conclusion is that the inflation fitted value term is not significant in the stock return equation, therefore inflation can be considered exogenous for stock returns. The same happens in inflation equation: fitted stock return term is not significant in the inflation equation, suggesting that stock returns are exogenous.

4 VAR & Cointegration with Trend

```
"""
@author: dtoppo
VAR & Cointegration with Trend
"""

import pandas as pd
import numpy as np
from statsmodels.formula.api import ols
from statsmodels.tsa.vector_ar import var_model

# Import dataserie from Excel
dataFile = "data_python_HW_2017.xlsx"
xlsx = pd.ExcelFile(dataFile)

sheetName = "VAR & Cointegration trend"
dataSeries = xlsx.parse(sheetName)

# =====
# Stationarity testing
# =====

diff = dataSeries.diff()
lag = dataSeries.shift()
difflag = diff.shift()

dependentVar = "WTF"

trend = pd.DataFrame(np.arange(0, np.size(dataSeries[dependentVar]), 1),
                    index=dataSeries.index)
trend = trend.shift(1)

data = pd.concat([diff[dependentVar], lag[dependentVar],
                difflag[dependentVar], trend], axis=1)
data.columns=["diff", "lag", "difflag", "trend"]
model = ols("diff ~ lag + difflag + trend", data[2:]).fit()
print(model.summary())
```


OLS Regression Results						
Dep. Variable:	diff	R-squared:	0.080			
Model:	OLS	Adj. R-squared:	0.073			
Method:	Least Squares	F-statistic:	11.28			
Date:	Thu, 25 May 2017	Prob (F-statistic):	4.15e-07			
Time:	15:31:15	Log-Likelihood:	-1139.9			
No. Observations:	393	AIC:	2288.			
Df Residuals:	389	BIC:	2304.			
Df Model:	3					
Covariance Type:	nonrobust					
	coef	std err	t	P> t	[0.025	0.975]
Intercept	0.1770	0.448	0.395	0.693	-0.703	1.057
lag	-0.0336	0.012	-2.801	0.005	-0.057	-0.010
difflag	0.2667	0.049	5.426	0.000	0.170	0.363
trend	0.0063	0.003	1.995	0.047	9.17e-05	0.012
Omnibus:	71.155	Durbin-Watson:	2.039			
Prob(Omnibus):	0.000	Jarque-Bera (JB):	556.260			
Skew:	-0.487	Prob(JB):	1.62e-121			
Kurtosis:	8.747	Cond. No.	466.			

Figure 12: Crude Oil WTI

```

dependentVar = "USDX"
data = pd.concat([diff[dependentVar], lag[dependentVar],
                 difflag[dependentVar], trend], axis=1)
data.columns=["diff", "lag", "difflag", "trend"]
model = ols("diff ~ lag + difflag + trend", data[2:]).fit()
print(model.summary())

```

OLS Regression Results						
Dep. Variable:	diff	R-squared:	0.010			
Model:	OLS	Adj. R-squared:	0.002			
Method:	Least Squares	F-statistic:	1.281			
Date:	Thu, 25 May 2017	Prob (F-statistic):	0.281			
Time:	15:32:25	Log-Likelihood:	-921.17			
No. Observations:	393	AIC:	1850.			
Df Residuals:	389	BIC:	1866.			
Df Model:	3					
Covariance Type:	nonrobust					
	coef	std err	t	P> t	[0.025	0.975]
Intercept	1.4798	1.121	1.320	0.188	-0.724	3.684
lag	-0.0158	0.010	-1.630	0.104	-0.035	0.003
diffflag	0.0095	0.051	0.186	0.852	-0.090	0.109
trend	-0.0002	0.001	-0.116	0.908	-0.003	0.003
Omnibus:	9.892	Durbin-Watson:	2.001			
Prob(Omnibus):	0.007	Jarque-Bera (JB):	17.972			
Skew:	0.007	Prob(JB):	0.000125			
Kurtosis:	4.048	Cond. No.	2.11e+03			

Figure 13: The U.S. Dollar Index(USDx)

Because for both prices and dividends, the ADF t Stat fails to exceed any ADF critical value (-3.42 for 5% confidence interval), we fail to reject the null hypothesis of non-stationarity of the two series.

Please note that the usual p-value (next to t Stat) doesn't make any sense here. It is the value if we have t distribution. Under the unit root process, we have Tau distribution. So, we have to compare computed t Stat with the critical values (given before) which were computed using the Tau distribution.

```

# Stationarity of difference
#
diffdiff = diff.diff()
diffdifflag = diffdiff.shift()

dependentVar = "WTF"

trend = pd.DataFrame(np.arange(0, np.size(dataSeries[dependentVar]), 1),
                    index=dataSeries.index)
trend = trend.shift(2)

data = pd.concat([diffdiff[dependentVar], diffflag[dependentVar],
                diffdifflag[dependentVar], trend], axis=1)
data.columns=["diffdiff", "diffflag", "diffdifflag", "trend"]
model = ols("diffdiff ~ diffflag + diffdifflag + trend", data[3:]).fit()

```

```
print(model.summary())
```

OLS Regression Results						
Dep. Variable:	diffdiff	R-squared:	0.378			
Model:	OLS	Adj. R-squared:	0.373			
Method:	Least Squares	F-statistic:	78.47			
Date:	Thu, 25 May 2017	Prob (F-statistic):	1.08e-39			
Time:	15:33:16	Log-Likelihood:	-1140.9			
No. Observations:	392	AIC:	2290.			
Df Residuals:	388	BIC:	2306.			
Df Model:	3					
Covariance Type:	nonrobust					
	coef	std err	t	P> t	[0.025	0.975]
Intercept	0.1219	0.452	0.270	0.788	-0.767	1.011
diffflag	-0.7134	0.062	-11.449	0.000	-0.836	-0.591
diffdiffflag	-0.0518	0.051	-1.020	0.308	-0.152	0.048
trend	-0.0006	0.002	-0.303	0.762	-0.005	0.003
Omnibus:	100.824	Durbin-Watson:	1.998			
Prob(Omnibus):	0.000	Jarque-Bera (JB):	757.230			
Skew:	-0.861	Prob(JB):	3.71e-165			
Kurtosis:	9.588	Cond. No.	454.			

Figure 14: Crude Oil WTI

```
dependentVar = "USDX"
data = pd.concat([diffdiff[dependentVar], diffflag[dependentVar],
                 diffdiffflag[dependentVar], trend], axis=1)
data.columns=["diffdiff", "diffflag", "diffdiffflag", "trend"]
model = ols("diffdiff ~ diffflag + diffdiffflag + trend", data[3:]).fit()
print(model.summary())
```

OLS Regression Results						
Dep. Variable:	diffdiff	R-squared:	0.506			
Model:	OLS	Adj. R-squared:	0.502			
Method:	Least Squares	F-statistic:	132.6			
Date:	Thu, 25 May 2017	Prob (F-statistic):	3.92e-59			
Time:	15:33:59	Log-Likelihood:	-918.02			
No. Observations:	392	AIC:	1844.			
Df Residuals:	388	BIC:	1860.			
Df Model:	3					
Covariance Type:	nonrobust					
	coef	std err	t	P> t	[0.025	0.975]
Intercept	-0.2825	0.257	-1.100	0.272	-0.788	0.223
diff1lag	-0.8894	0.071	-12.473	0.000	-1.030	-0.749
diff2lag	-0.1105	0.050	-2.191	0.029	-0.210	-0.011
trend	0.0012	0.001	1.020	0.308	-0.001	0.003
Omnibus:	13.637	Durbin-Watson:	2.011			
Prob(Omnibus):	0.001	Jarque-Bera (JB):	28.248			
Skew:	-0.094	Prob(JB):	7.34e-07			
Kurtosis:	4.302	Cond. No.	456.			

Figure 15: The U.S. Dollar Index(USDx)

For both Oil and USDx, the ADF t Stat exceeds the ADF critical value (5%, -3.42), we reject the null hypothesis of non-stationarity; we then conclude that Oil and USDx are I(1).

```

#-----
# OLS Oil vs USDx
# The - sign can be used to remove columns/variables. For instance,
# we can remove the intercept from a model by adding "-1" to the formula
#-----
model = ols("WTI ~ USDx -1", dataSeries).fit()
print(model.summary())

```

OLS Regression Results						
Dep. Variable:	WTI	R-squared:	0.567			
Model:	OLS	Adj. R-squared:	0.566			
Method:	Least Squares	F-statistic:	516.3			
Date:	Thu, 25 May 2017	Prob (F-statistic):	1.19e-73			
Time:	15:34:45	Log-Likelihood:	-1950.2			
No. Observations:	395	AIC:	3902.			
Df Residuals:	394	BIC:	3906.			
Df Model:	1					
Covariance Type:	nonrobust					
	coef	std err	t	P> t	[0.025	0.975]
USDX	0.3991	0.018	22.723	0.000	0.365	0.434
Omnibus:	52.586	Durbin-Watson:	0.021			
Prob(Omnibus):	0.000	Jarque-Bera (JB):	71.087			
Skew:	1.030	Prob(JB):	3.66e-16			
Kurtosis:	2.732	Cond. No.	1.00			

Figure 16: OLS for Cointegration

```

#-----
# Lagged cointegration
#-----
residuals = model.wresid.to_frame()
diffResiduals = residuals.diff()
laggedResiduals = residuals.shift()

diffResiduals.columns = ["diffResiduals"]
laggedResiduals.columns = ["laggedResiduals"]

data = pd.concat([diffResiduals, laggedResiduals], axis=1)
model = ols("diffResiduals ~ laggedResiduals -1", data[1:]).fit()
print(model.summary())

```

OLS Regression Results						
Dep. Variable:	diffResiduals	R-squared:	0.006			
Model:	OLS	Adj. R-squared:	0.003			
Method:	Least Squares	F-statistic:	2.249			
Date:	Thu, 25 May 2017	Prob (F-statistic):	0.134			
Time:	15:44:25	Log-Likelihood:	-1185.6			
No. Observations:	394	AIC:	2373.			
Df Residuals:	393	BIC:	2377.			
Df Model:	1					
Covariance Type:	nonrobust					
	coef	std err	t	P> t	[0.025	0.975]
laggedResiduals	-0.0110	0.007	-1.500	0.134	-0.025	0.003
Omnibus:	106.472	Durbin-Watson:	1.524			
Prob(Omnibus):	0.000	Jarque-Bera (JB):	1018.590			
Skew:	-0.839	Prob(JB):	6.55e-222			
Kurtosis:	10.696	Cond. No.	1.00			

Figure 17: Summary Output

The t Stat is 1.49. The 5% critical value for the model with no intercept is 3.37.

The t Stat falls within the non-rejection region and the null hypothesis of no cointegration is NOT rejected.

Of course this is not Johansen's method and it appears to be primitive when compared to it.

```

#
# VAR(2)
# First using regression
#
diffflaglag = diffflag.shift()

data = pd.concat([diff["WTI"], diffflag["WTI"], diffflaglag["WTI"],
diffflag["USDX"], diffflaglag["USDX"]], axis=1)
data.columns=["WTI", "diffflagWTI", "diffflaglagWTI", "diffflagUSDX", "diffflaglagUSDX"]
model = ols("WTI ~ diffflagWTI + diffflaglagWTI + diffflagUSDX + diffflaglagUSDX",
data[3:]).fit()
print(model.summary())

```

OLS Regression Results

Dep. Variable:	WTI	R-squared:	0.069
Model:	OLS	Adj. R-squared:	0.060
Method:	Least Squares	F-statistic:	7.189
Date:	Thu, 25 May 2017	Prob (F-statistic):	1.36e-05
Time:	15:45:54	Log-Likelihood:	-1139.8
No. Observations:	392	AIC:	2290.
Df Residuals:	387	BIC:	2309.
Df Model:	4		
Covariance Type:	nonrobust		

	coef	std err	t	P> t	[0.025	0.975]
Intercept	0.0019	0.225	0.008	0.993	-0.441	0.445
difflagWTI	0.2475	0.052	4.761	0.000	0.145	0.350
difflaglagWTI	0.0377	0.052	0.722	0.471	-0.065	0.140
difflagUSDX	0.0786	0.091	0.860	0.390	-0.101	0.258
difflaglagUSDX	-0.1113	0.091	-1.217	0.224	-0.291	0.069

Omnibus:	102.077	Durbin-Watson:	2.000
Prob(Omnibus):	0.000	Jarque-Bera (JB):	724.669
Skew:	-0.894	Prob(JB):	4.37e-158
Kurtosis:	9.416	Cond. No.	5.18

Figure 18: Crude Oil WTI

```
data = pd.concat([diff["USDX"], difflag["WTI"], difflaglag["WTI"],
                 difflag["USDX"], difflaglag["USDX"]], axis=1)
data.columns=["USDX", "difflagWTI", "difflaglagWTI", "difflagUSDX", "difflaglagUSDX"]
model = ols("USDX ~ difflagWTI + difflaglagWTI + difflagUSDX + difflaglagUSDX",
            data[3:]).fit()
print(model.summary())
```

OLS Regression Results						
Dep. Variable:	USDX	R-squared:	0.021			
Model:	OLS	Adj. R-squared:	0.011			
Method:	Least Squares	F-statistic:	2.041			
Date:	Thu, 25 May 2017	Prob (F-statistic):	0.0880			
Time:	15:48:31	Log-Likelihood:	-916.99			
No. Observations:	392	AIC:	1844.			
Df Residuals:	387	BIC:	1864.			
Df Model:	4					
Covariance Type:	nonrobust					
	coef	std err	t	P> t	[0.025	0.975]
Intercept	-0.0557	0.128	-0.436	0.663	-0.307	0.195
diff1lagWTI	-0.0503	0.029	-1.708	0.088	-0.108	0.008
diff2lagWTI	0.0236	0.030	0.798	0.425	-0.035	0.082
diff1lagUSDX	-0.0152	0.052	-0.293	0.769	-0.117	0.087
diff2lagUSDX	0.1218	0.052	2.351	0.019	0.020	0.224
Omnibus:	15.691	Durbin-Watson:	2.004			
Prob(Omnibus):	0.000	Jarque-Bera (JB):	31.726			
Skew:	-0.175	Prob(JB):	1.29e-07			
Kurtosis:	4.349	Cond. No.	5.18			

Figure 19: The U.S. Dollar Index(USDX)

The series are not cointegrated, therefore we estimate a VAR(2) on their first differences (since both the series are non-stationary in levels).

Using a 5% significance level, DiffOil is significantly related to its own past values, while DiffUSDX is significantly related to Lag2DiffUDSX.

5 GARCH

```
"""
@author: dtoppo
GARCH
"""

import numpy as np
from scipy import optimize
import time
import pandas as pd

# Import dataserie from Excel
dataFile = "data_python_HW_2017.xlsx"
xlsx = pd.ExcelFile(dataFile)

sheetName = "GARCH"
dataSeries = xlsx.parse(sheetName)

=====
# MAX Likelihood function
#
def minLikelihood(params, data):
    s = GARCH(params, data)
    logL = -((-1/2 * np.log(2*np.pi) - 1/2 * np.log(s) - 1/2 * data**2/s).sum())
    return logL

=====
# MAX Likelihood function for Variance Targeting
#
def minLikelihoodVarTarget(params, data):
    s = GARCHTargetVar(params, data)
    logL = -((-1/2 * np.log(2*np.pi) - 1/2 * np.log(s) - 1/2 * data**2/s).sum())
    return logL

=====
# GARCH(1,1)
#
def GARCH(param, data):
    alpha, beta, omega = param
    s = np.zeros(len(data))
    s[0] = np.var(data, ddof=1)

    for i in range(1, len(data)):
        s[i] = omega + alpha*data[i-1]**2 + beta*(s[i-1]) # GARCH(1,1) model
    return s

=====
# GARCH(1,1) with Variance Targeting
#
def GARCHTargetVar(param, data):
    alpha, beta, omega = param
    s = np.zeros(len(data))
    s[0] = np.var(data, ddof=1)
    omega=s[0]*(1-alpha-beta)

    for i in range(1, len(data)):
```

```

        s[i] = omega + alpha*data[i-1]**2 + beta*(s[i-1]) # GARCH(1,1) model
    return s

#=====
# Optimizer
#=====
def maximizeMLE(initParams, data):
    # Constraints
    # 1 - (alpha*(1 + theta^2)+beta) >= 0
    def persistenceIndexConstraint(params):
        omega, alpha, beta = params
        return 1 - (alpha + beta)

    cons = {'type' : 'ineq', 'fun' : persistenceIndexConstraint}

    # Could be also a lambda expression
    # cons = {'type' : 'ineq', 'fun' : lambda params : 1 - np.sum(params)}

    # Bounds
    # 0 <= parameters <= 1
    bnds = ((0, 1), (0, 1), (0, 1))

    # Run the minimizer
    results = optimize.minimize(minLikelihood, initParams, data, method='SLSQP',
                                bounds=bnds, constraints=cons)

    return results.x

#=====
# Optimizer Variance Targeting
#=====
def maximizeMLEVarTarget(initParams, data):
    # Constraints
    # 1 - (alpha*(1 + theta^2)+beta) >= 0
    def persistenceIndexConstraint(params):
        omega, alpha, beta = params
        return 1 - (alpha + beta)

    cons = {'type' : 'ineq', 'fun' : persistenceIndexConstraint}

    # Bounds
    # 0 <= parameters <= 1
    bnds = ((0, 1), (0, 1), (0, 1))

    # Run the minimizer
    results = optimize.minimize(minLikelihoodVarTarget, initParams, data, method='SLSQP',
                                bounds=bnds, constraints=cons)

    return results.x

#=====
# Main routine
#=====
def main() :

    data = dataSeries["Germany"]

```

```

# Initial parameter guesses (alpha, beta, omega)
initParams = [0.0832, 0.8759, 0.0001]

# GARCH
start = time.time()
results = maximizeMLE(initParams, data)
end = time.time()

# Print the results
print("Alpha: %.6f\nBeta: %.6f\nOmega: %.6f" % (results[0], results[1], results[2]))
print("Computed in %.2f secs" % (end - start))
print()

```

Alpha		0.1480
Beta		0.8040
Omega		0.0002

The persistence index ($0.148 + 0.804 = 0.952$) is lower than 1. A high persistence implies that shocks, which may push variance away from its long-run average, will persist for a long time.

```

# GARCH with Variance Targeting
start = time.time()
results = maximizeMLEVarTarget(initParams, data)
end = time.time()

# Print the results
print("Alpha: %.6f\nBeta: %.6f\nOmega: %.6f" % (results[0], results[1], results[2]))
print("Computed in %.2f secs" % (end - start))

```

Alpha		0.1374
Beta		0.8062
Omega		0.0002

6 GARCH with Leverage

```
"""
@author: dtoppo
GARCH with Leverage
"""

import pandas as pd
import numpy as np
from scipy import optimize
import matplotlib.pyplot as plt
from statsmodels.formula.api import ols

# Import dataserie from Excel
dataFile = "data_python-HW_2017.xlsx"
xlsx = pd.ExcelFile(dataFile)

sheetName = "GARCH with Leverage"
dataSeries = xlsx.parse(sheetName)

#=====#
# MAX Likelihood function
#=====#
def minLikelihood(params, data):
    s = GARCHLeveraged(params, data)
    logL = -((-1/2 * np.log(2*np.pi) - 1/2 * np.log(s) - 1/2 * data**2/s).sum())
    return logL

#=====#
# GARCH(1,1) with leverage model
#=====#
def GARCHLeveraged(params, data):
    alpha, beta, omega, theta = params

    s = np.zeros(np.size(data))
    s[0] = np.var(data, ddof=1)

    for i in range(1, np.size(data)):
        s[i] = omega + alpha*((data[i-1] - theta * np.sqrt(s[i-1]))**2) + beta*(s[i-1])
    return s

#=====#
# Main routine
#=====#
def maximizeMLE(initParams, data):
    alpha, beta, omega, theta = initParams

    # Constraints
    # 1 - (alpha*(1 + theta^2)+beta) >= 0
    def persistenceIndexConstraint(params):
        return 1 - (alpha*(1 + theta**2) + beta)

    cons = {'type' : 'ineq', 'fun' : persistenceIndexConstraint}

    # Could be also a lambda expression
    # cons = {'type' : 'ineq',
    #         'fun' : lambda params : 1 - (params[0]*(1 + params[2]**2) + params[1])}
```

```

# Bounds
# 0 <= parameters <= 1
bnds = ((0, 1), (0, 1), (0, 1), (0, 1))

# Run the minimizer - SLSQP
# SLSQP uses Sequential Least Squares Programming to minimize a
# function of several variables
# with any combination of bounds, equality and inequality constraints.
results = optimize.minimize(minLikelihood, initParams, data,
                             method='SLSQP', bounds=bnds, constraints=cons)
return results.x

=====
# Lagged autocorrelation function
#
def laggedSquaredAutoCorrelation(data, nbLags):

    rSquared = data**2

    autocorrel = np.zeros(nbLags)
    baseSquaredArray = pd.DataFrame(rSquared[0: np.size(rSquared) - 1])

    for i in np.arange(0, nbLags, 1):
        lagged = baseSquaredArray.shift(-i-1).fillna(0)
        autocorrel[i] = np.corrcoef(baseSquaredArray[0].values, lagged[0].values)[0,1]

    return autocorrel

=====
# Squared Standard auto correlation
#
def laggedSquaredStandardAutoCorrelation(data, nbLags):

    rSquared = data**2

    autocorrel = np.zeros(nbLags)

    # Initial parameter guesses (alpha, beta, omega, theta)
    initParams = [0.2, 0.8, 0.0, 0.0]

    garchParams = maximizeMLE(initParams, data)
    garchVol = GARCHLeveraged(garchParams, data)

    squaredStandardizedRet = rSquared / garchVol

    baseSquaredArray = squaredStandardizedRet[0: np.size(squaredStandardizedRet) - 1]
    lagged = baseSquaredArray

    for i in np.arange(0, nbLags, 1):
        lagged = np.delete(lagged, 0)
        lagged = np.append(lagged, 0)
        autocorrel[i] = np.corrcoef(baseSquaredArray, lagged)[0,1]

    return autocorrel

=====

```

```

# Main routine
#
def main():

    seriesName = "Italy"

    # Initial parameter guesses (alpha, beta, omega, theta)
    initParams = [0.1, 0.85, 0.000005, 0.0]

    results = maximizeMLE(initParams, dataSeries[seriesName].values)

    # Print the results
    print("\nAlpha: %.6f\nBeta: %.6f\nOmega: %.6f\nTheta: %.6f" % (results[0], results[1],
        results[2], results[3]))

```

Alpha	0.2069
Beta	0.5978
Omega	0.0010
Theta	0.1129

Leverage Effect: a negative return on a stock implies a drop in the equity value, which implies that the company becomes more highly levered and thus riskier (assuming the level of debt stays constant).

```

#
# Plot sample autocorrelation coefficients at lags 1 through 100
#
nbLags = 100

squaredAutocorrel = laggedSquaredAutoCorrelation(dataSeries[seriesName].values, nbLags)
squaredStandardAutoCorrel =
    laggedSquaredStandardAutoCorrelation(dataSeries[seriesName].values, nbLags)
lag = np.arange(1, nbLags + 1, 1)

barlettLowerBand = np.array(np.ones(100))
barlettLowerBand = barlettLowerBand * (-1.96 / np.size(dataSeries[seriesName])**0.5)
barlettUpperBand = -barlettLowerBand

# Plot data
plt.plot(lag, squaredAutocorrel)
plt.plot(lag, squaredStandardAutoCorrel)

# Plot bands
plt.plot(barlettLowerBand, '—r')
plt.plot(barlettUpperBand, '—r')

```

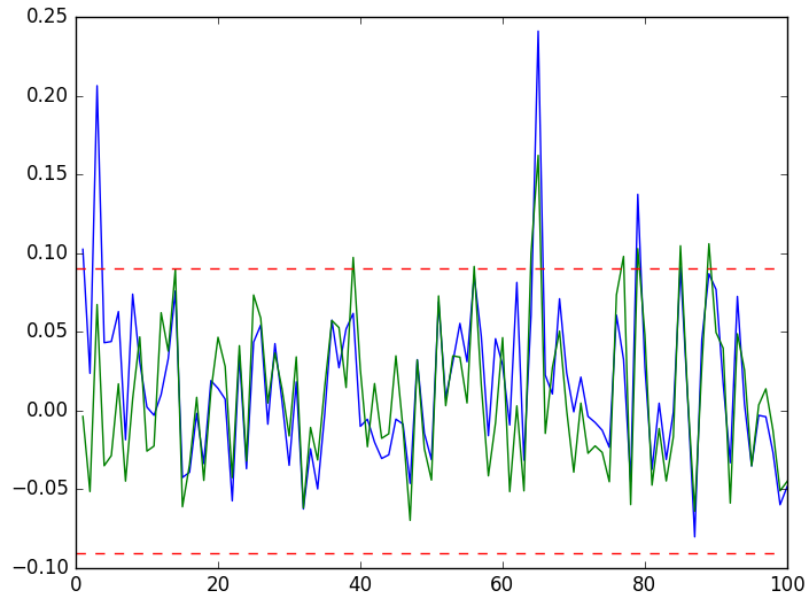


Figure 20: Sample Autocorrelation

```

#
# Regression of daily squared returns on variance
#
forecastedVar = GARCHLeveraged(results , dataSet[seriesName].values)
squaredReturns = dataSet[seriesName].values**2

sqretDataFrame = pd.DataFrame(data = squaredReturns ,
                              index = dataSet.index , columns = ["sqRet"])
garchDataFrame = pd.DataFrame(data = forecastedVar ,
                              index = dataSet.index , columns = ["garchVar"])

data = pd.concat([sqretDataFrame , garchDataFrame] , axis=1)
model = ols("sqRet ~ garchVar" , data).fit()
print(model.summary())

```

OLS Regression Results						
Dep. Variable:	sqRet	R-squared:	0.021			
Model:	OLS	Adj. R-squared:	0.019			
Method:	Least Squares	F-statistic:	10.17			
Date:	Thu, 25 May 2017	Prob (F-statistic):	0.00152			
Time:	16:33:22	Log-Likelihood:	1524.6			
No. Observations:	468	AIC:	-3045.			
Df Residuals:	466	BIC:	-3037.			
Df Model:	1					
Covariance Type:	nonrobust					
	coef	std err	t	P> t	[0.025	0.975]
Intercept	0.0023	0.001	2.343	0.020	0.000	0.004
garchVar	0.5351	0.168	3.189	0.002	0.205	0.865
Omnibus:	458.729	Durbin-Watson:	2.032			
Prob(Omnibus):	0.000	Jarque-Bera (JB):	14357.875			
Skew:	4.393	Prob(JB):	0.00			
Kurtosis:	28.673	Cond. No.	389.			

Figure 21: Summary Output

The intercept is significantly different from zero at a 5% confidence level; the slope is around 0.54 and this estimate is significantly different from 1. Therefore the GARCH(1,1) model with leverage offers a poor variance model.

7 QQ Plots

```
"""
@author: dtoppo
QQ Plots
"""

import numpy as np
import matplotlib.pyplot as plt
import scipy as sp
import pandas as pd

from GARCHLeverage import maximizeMLE, GARCHLeveraged

# Import dataserie from Excel
dataFile = "data_python_HW_2017.xlsx"
xlsx = pd.ExcelFile(dataFile)

sheetName = "QQ Plot"
dataSeries = xlsx.parse(sheetName)

#=====#
# Get descriptive statistics
#=====#
def descriptiveStats(data):
    nbObs = len(data)
    mean = np.mean(data)
    stdev = np.std(data, ddof=1)
    skewness = sp.stats.skew(data, bias=False)
    kurtosis = sp.stats.kurtosis(data, bias=False)

    result = {"nbObs" : nbObs, "mean" : mean, "stdev" : stdev,
              "skewness" : skewness, "kurtosis" : kurtosis}
    return result

#=====#
# Descriptive statistics printing helper
#=====#
def printDescriptiveStats(stats):
    print("Descriptive Statistics")
    print("Number of obs.\t%d" %stats["nbObs"])
    print("Mean\t\t%.4f%%" %(stats["mean"] * 100))
    print("Std. Deviation\t%.4f%%" %(stats["stdev"] * 100))
    print("Skewness\t%.4f" %stats["skewness"])
    print("Kurtosis\t%.4f" %stats["kurtosis"])

#=====#
# QQ Plot function
#=====#
def qqPlot(normalizedReturns):

    nbObs = len(normalizedReturns)

    sortedNormalizedReturns = np.sort(normalizedReturns)
    normalizedQuantiles = sp.stats.norm.ppf((np.arange(1, nbObs+1) - 0.5)/nbObs)
```

```

plt.figure()
plt.plot(normalizedQuantiles, sortedNormalizedReturns, 'rD')

# Establish min and max values for axes
min_value = np.around(sortedNormalizedReturns[0])
max_value = np.around(sortedNormalizedReturns[len(sortedNormalizedReturns) - 1])

absolute_value = max((np.abs(min_value), max_value))

plt.xlim(-absolute_value, absolute_value)
plt.ylim(-absolute_value, absolute_value)

# Bisectrix plot
x = np.arange(-absolute_value, absolute_value + 1, 1)
plt.plot(x, x)

def main():

    #=====
    # Simple QQ Plot
    #=====
    data = dataSeries["Italy"].values

    # Standard statistics
    stats = descriptiveStats(data)
    printDescriptiveStats(stats)

    Number of obs. | 468
                   | Mean 1.0732%
    Std. Deviation | 6.9915%
                   | Skewness 0.5249
                   | Kurtosis 1.3336

    # normalize returns
    normalizedReturns = data / stats["stdev"];

    # QQ Plot
    qqPlot(normalizedReturns)

```

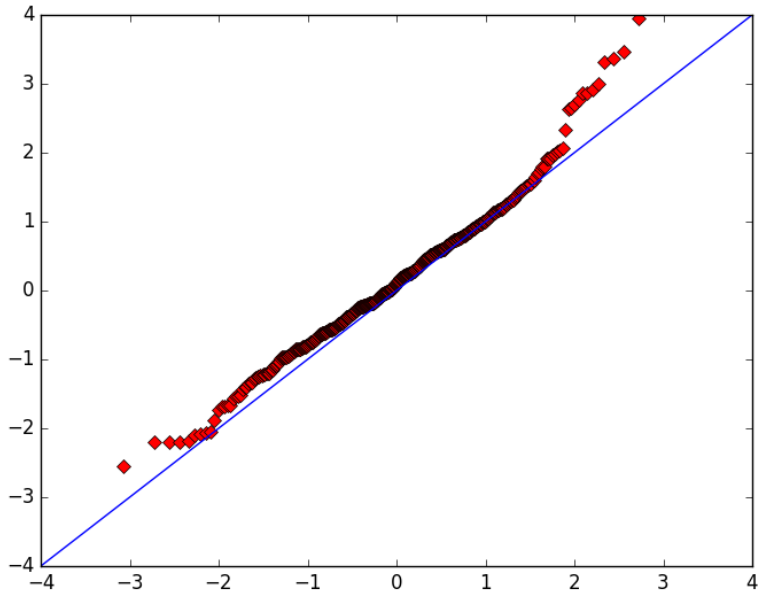


Figure 22: QQ Plot of Italian Equity Returns (scaled by unconditional variance)

```

#-----
# QQ Plot using GARCH vol
#-----
data = dataSeries["Italy"].values

# GARCH vol
initParams = [0.2, 0.8, 0.0, 0.0]
garchParams = maximizeMLE(initParams, data)
garchVol = GARCHLeveraged(garchParams, data)

# normalize returns
normalizedReturns = data / np.sqrt(garchVol)

# Standard statistics
stats = descriptiveStats(normalizedReturns)
printDescriptiveStats(stats)

```

Number of obs.	468
Mean	15.0207%
Std. Deviation	99.2437%
Skewness	0.3608
Kurtosis	0.8175

```

# QQ Plot
qqPlot(normalizedReturns)

```

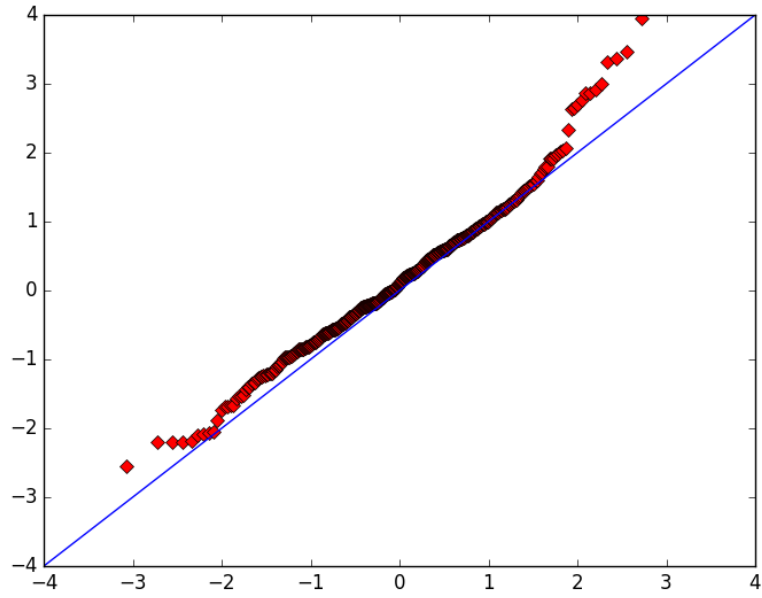


Figure 23: QQ Plot of Italian Equity Returns with GARCH(1,1) Shocks

8 Correlations

```
"""
@author: dtoppo
Correlations
"""

import numpy as np
import pandas as pd

# Import dataserie from Excel
dataFile = "data_python-HW_2017.xlsx"
xlsx = pd.ExcelFile(dataFile)

sheetName = "Correlation"

dataSeries = xlsx.parse(sheetName)

# Covariance matrix
cov_matrix = np.cov(dataSeries.T)
print("Covariance matrix: ")
print(cov_matrix)

      [0.00488812  0.00157657  0.00205688  0.00175714  0.00186691]
      [0.00157657  0.00201506  0.00176448  0.0014482  0.00159337]
      [0.00205688  0.00176448  0.00291552  0.00170511  0.00160591]
      [0.00175714  0.0014482  0.00170511  0.00240827  0.00154439]
      [0.00186691  0.00159337  0.00160591  0.00154439  0.00304086]

# Determinant
determinant = np.linalg.det(cov_matrix)
print("Determinant : %.6f" % determinant)

Determinant | 0

# Correlation matrix
corr_matrix = np.corrcoef(dataSeries.T);
print("\n\nCorrelation matrix: ")
print(corr_matrix)

      [ 1.         0.50234023  0.54485513  0.51213205  0.48423435]
      [0.50234023  1.         0.72797427  0.65740331  0.64368594]
      [0.54485513  0.72797427  1.         0.64349111  0.53934197]
      [0.51213205  0.65740331  0.64349111  1.         0.57069908]
      [0.48423435  0.64368594  0.53934197  0.57069908  1.         ]

# Determinant
determinant = np.linalg.det(corr_matrix)
print("Determinant : %.6f" % determinant)

Determinant | 0.82543
```

9 Dynamic Conditional Correlations

```

"""
@author: dtoppo
Dynamic Conditional Correlations
"""

import pandas as pd
import numpy as np
import scipy as sp
from scipy import optimize
from scipy.special import gammaln

from GARCHLeverage import GARCHLeveraged, maximizeMLE

# Import dataserie from Excel
dataFile = "data_python_HW_2017.xlsx"
xlsx = pd.ExcelFile(dataFile)

sheetName = "DCC"
dataSeries = xlsx.parse(sheetName)

# =====
# Unconditional sample covariance and correlation
# =====
sampleCov = np.cov(dataSeries[["Germany", "Japan"]].T, ddof=0)[0,1]
sampleCorr = np.corrcoef(dataSeries[["Germany", "Japan"]].T, ddof=0)[0,1]

# =====
# Portfolio unconditional covariance
# =====
GermanyWgt = 0.5;
JapanWgt = 0.5;

GermanyVar = np.var(dataSeries["Germany"], ddof=1)
JapanVar = np.var(dataSeries["Japan"], ddof=1)

# PPF = percent point function == quantile function = norminv
# q = prob(X<=x)
VaRGermany = -sp.stats.norm.ppf(q=0.01, loc=0, scale=1)*np.sqrt(GermanyVar)*GermanyWgt
VaRJapan = -sp.stats.norm.ppf(q=0.01, loc=0, scale=1)*np.sqrt(JapanVar)*JapanWgt
sumOfVaRs = VaRGermany + VaRJapan

pfCov = np.cov(dataSeries[["Germany", "Japan"]].T, ddof=0)[0,1]
pfVar = GermanyWgt**2*GermanyVar + JapanWgt**2*JapanVar + 2*GermanyWgt*JapanWgt*pfCov

VaRPf = -sp.stats.norm.ppf(q=0.01, loc=0, scale=1)*np.sqrt(pfVar)

print(sumOfVaRs);
print(VaRPf);

Sum of VaRs | 0.1239
Portfolio VaRs | 0.1035

```

Due to diversification, the VaR of the portfolio is less than the sum of the individual VaRs.

```

#=====
# NAGARCH(1,1)
#=====
deData = dataSeries["Germany"].values

deInitParams = [0.2, 0.5, 0.0008, 0.75]
deParams = maximizeMLE(deInitParams, deData)
deGarchVol = GARCHLeveraged(deParams, deData)
deStdRet = deData / np.sqrt(deGarchVol)

jpData = dataSeries["Japan"].values

jpInitParams = [0.2, 0.5, 0.0008, 0.75]
jpParams = maximizeMLE(jpInitParams, jpData)
jpGarchVol = GARCHLeveraged(jpParams, jpData)
jpStdRet = jpData / np.sqrt(jpGarchVol)

#=====
# Lamda estimation
#=====
def computeQ(lda):
    qDE = np.ones(np.size(deData))
    qDEJP = np.ones(np.size(deData))
    qJP = np.ones(np.size(deData))

    q = pd.DataFrame(data=np.array([qDE, qDEJP, qJP]).T,
                     columns=["DE-DE", "DE-JP", "JP-JP"], index=dataSeries.index)

    q["DE-DE"][0] = 1
    q["DE-JP"][0] = (deStdRet * jpStdRet).sum() / np.size(deStdRet)
    q["JP-JP"][0] = 1

    for t in range(1, np.size(deData)):
        q["DE-DE"][t] = (1-lda)*deStdRet[t-1]**2 + lda*q["DE-DE"][t-1]
        q["DE-JP"][t] = (1-lda)*deStdRet[t-1]*jpStdRet[t-1] + lda*q["DE-JP"][t-1]
        q["JP-JP"][t] = (1-lda)*jpStdRet[t-1]**2 + lda*q["JP-JP"][t-1]

    return q

def dccMinLikelihood(lda):
    q = computeQ(lda)
    r = q["DE-JP"] / np.sqrt(q["DE-DE"] * q["JP-JP"])
    logL = -1/2*(np.log(1-r**2)) + (deStdRet**2+jpStdRet**2-2*r*deStdRet*jpStdRet)/(1-r**2)
    return -(logL.sum())

bnds = [(0.00001, 0.99999)]
initParam = 0.94
resultDcc = optimize.minimize(dccMinLikelihood, initParam,
                              method='SLSQP', bounds=bnds)

print(resultDcc)

Optimal Lambda | 0.9841

#=====
# GARCH(1,1)-t(d) model
#=====

```

```

#
data = dataSeries["Italy"]

# Initialize GARCH(1,1)
def garch(params):
    omega, alpha, beta, delta, d = params
    s = np.zeros(np.size(data))
    s[0] = np.var(data, ddof=1)
    for i in range(1, np.size(data)):
        s[i] = omega + alpha*(data[i-1] - delta*np.sqrt(s[i-1]))**2 + beta*s[i-1]
    return s

# MLE
def mle(params):
    s = garch(params)
    di = params[4]
    logL = (gammaln(di*0.5 + 0.5) - 0.5*np.log(np.pi) - gammaln(di*0.5) - \
            0.5*np.log(di-2)-(di*0.5 + 0.5)*np.log(1+(data**2/s)/(di - 2)) \
            - 0.5*np.log(s)).sum()
    return -logL

# Optimization
initParameters = [0.0001, 0.085, 0.876, 0.5, 3]
bnds = [(0.0001,0.9999), (0.0001,0.9999), (0.0001,0.9999), (0.0001,None),
        (2.0001,None)]

def persistence(params):
    omega, alpha, beta, delta, d = params
    return 1 - (alpha*(1 + delta**2) + beta)

cons = {"type": "ineq", "fun": persistence}

result = optimize.minimize(mle, initParameters, method="SLSQP", bounds=bnds,
                           constraints=cons)

print(result)

```

Alpha	0.1373
Beta	0.7616
Omega	0.0005
Theta	0.2041
d	7.5876