# ON DEFINING NEIGHBOURHOODS OF MEASURES THROUGH THE CONCENTRATION FUNCTION

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SUMMARY. Statistical procedures are often interested in comparing probability measures by means of distances defined over the space  $\mathcal{P}$  of all probability measures, endowed with some classical topology, like the variational or the Prohorov ones. Other topologies can be obtained by means of the concentration function, which extends the notion of Lorenz curve. Hence, neighbourhood classes  $\Gamma$  of probability measures, including well-known ones, are defined and a representation theorem is proved. Finally, ranges of functionals over  $\Gamma$  are found, restricting the search among the extremal measures in  $\Gamma$ .

## 1. Comparison of probability measures

Many statistical problems require the specification of topologies or distances on the space  $\mathcal{P}$  of all probability measures, e.g. to study similarities among populations or to consider neighbourhoods of a given probability measure. Suppose that we are interested in comparing the functional forms of two probability measures, say P and  $P_0$ , on the same measurable space  $(\Theta, \mathcal{F})$ ,  $\Theta$  being a Polish space and  $\mathcal{F}$  its Borel  $\sigma$ -field. We could use the variational distance  $d_V(P, P_0) = \sup_{A \in \mathcal{F}} |P(A) - P_0(A)|$  or the Prohorov distance  $d_P(P, P_0) = \inf\{\varepsilon > 0 : P(A) \leq P_0(A^{\varepsilon}) + \varepsilon \ \forall \ A \in \mathcal{F}\}$ , where  $A^{\varepsilon} = \{\theta \in \Theta : d(\theta, A) \leq \varepsilon\}$  and d is a metric on  $\Theta$ .

However, such rules do not seem sufficiently sensitive on the sets with small probability under  $P_0$ . For example, if the variational metric is considered, then a  $\varepsilon$ -neighbourhood of  $P_0$  contains all the probability measures P such that, for any  $A \in \mathcal{F}, |P(A)-P_0(A)| \leq \varepsilon$ . Consider now a set E such that  $P_0(E) = \varepsilon/10$ . Given P in the  $\varepsilon$ -neighbourhood of  $P_0$ , it follows that  $P(E) \leq 11\varepsilon/10$  is the only restriction about P on E; i.e. P is considered close to  $P_0$  even if its value on E is cleven times greater than  $P_0(E)$ . A similar reasoning holds for the  $\varepsilon$ -contamination class of probability measures, described in Huber (1981), which contains all the probability measures P such that, for any  $A \in \mathcal{F}, (1-\varepsilon)P_0(A) \leq P(A) \leq (1-\varepsilon)P_0(A)+\varepsilon$ . When such a consequence is deemed inconvenient, then different bounds on P(A) could be

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considered and the concentration function (c.f.) is a flexible tool to get them. As an example, require, for any  $A \in \mathcal{F}$ , either  $|P_0(A) - P(A)| \le \varepsilon \min\{P_0(A), 1 - P_0(A)\}$  or  $|P_0(A) - P(A)| \le P_0(A)(1 - P_0(A))$ , so that more stringent bounds are found on P(E); in the former case, we have  $\frac{\varepsilon}{10}(1 - \varepsilon) \le P(E) \le \frac{\varepsilon}{10}(1 + \varepsilon)$ , while the latter implies  $\frac{\varepsilon^2}{100} \le P(E) \le \frac{\varepsilon}{10}(2 - \frac{\varepsilon}{10})$ , i.e. P(E) does not exceed twice  $P_0(E)$ .

In this paper, we develop a method which enables us to define neighbourhoods of a probability measure, specifying bounds on the probability of any measurable subset  $A \in \mathcal{F}$ . Such a method is essentially based on the concentration function defined by Cifarelli and Regazzini (1987) as an extension of the classical notion of the Lorenz-Gini curve. A g-neighbourhood of a probability measure  $P_0$ , defined in Section 2, is made of all the probability measures whose concentration function with respect to  $P_0$  lies above a specified continuous, convex, monotone nondecreasing function g. In Section 3, it will be shown that g-neighbourhoods determine a topology over  $\mathcal{P}$ , while a representation theorem will be proved in Section 4. Computations of upper and lower bounds on functionals over g-neighbourhoods are simplified by the results in Section 5, while some final remarks are presented in Section 6.

### 2. Definition of g-neighbourhoods

In this section we consider classes  $K_g$  of probability measures which can be defined through the c.f.'s, as neighbourhoods around a base measure  $P_0$ .

Definition 1. If  $g:[0,1] \rightarrow [0,1]$  is a continuous, convex, monotone nondecreasing function such that g(0) = 0, then the set

$$K_{q} = \{ P \in \mathcal{P} : P(A) \ge g(P_{0}(A)) \ \forall \ A \in \mathcal{F} \}$$
 ... (1)

will be said a g-neighbourhood of  $P_0$ .

Observe that, if  $P \in K_g$ , then  $g(P_0(A)) \le P(A) \le 1 - g(1 - P_0(A))$ .

We give now the reasons for g to be continuous, monotone nondecreasing and convex. The requirement g(0) = 0 is obvious to get  $P(\emptyset) = 0$ .

Monotonicity. Let g(x) belong to the range of a measure  $P \in K_g$  and let  $A \in \mathcal{F}$  be such that  $P_0(A) = x$  and P(A) = g(x). If  $B \subset A$ , then  $P(B) \leq P(A)$ . Hence we choose a monotone nondecreasing function g.

Continuity. Let g(x) belong to the range of a measure  $P \in K_g$  and let  $A \in \mathcal{F}$  be such that  $P_0(A) = x$  and P(A) = g(x). Since P is a regular measure, it follows that  $P(A) = \inf_{\{G \text{ open}: A \subseteq G\}} P(G)$ . Therefore, for any  $\varepsilon > 0$  there exists an open set

G such that  $P(G) < P(A) + \varepsilon \le g(x) + \varepsilon$ . Since  $P(G) \ge g(P_0(G))$ , g must be right-continuous in order to be meaningful. Analogously, left-continuity is required.

Convexity. Let  $g(x_1)$  and  $g(x_2)$  belong to the range of a measure  $P \in K_g$  and suppose there exist  $A_1, A_2 \in \mathcal{F}$  such that  $A_1 \subset A_2, P_0(A_i) = x_i$  and  $P(A_i) = x_i$ 

 $g(x_i), i = 1, 2$ . If  $B_1$  and  $B_2$  partition  $A_2 \setminus A_1$ , then there exists  $B_i$  (i = 1 or 2) such that  $P_0(B_i) = \lambda P_0(A_2 \setminus A_1)$  and  $P(B_i) \leq \lambda P(A_2 \setminus A_1) \leq \lambda (g(x_2) - g(x_1))$ . Since

$$\inf_{P_0(C)\geq x_1+\lambda(x_2-x_1)} P(C) \leq P(A_1) + P(B_i) \leq g(x_1) + \lambda(g(x_2) - g(x_1)),$$

it follows that  $g((1 - \lambda)x_1 + \lambda x_2) \le (1 - \lambda)g(x_1) + \lambda g(x_2)$ . Hence the convexity is another reasonable requirement.

The definition of g-neighbourhood can be reformulated by means of the concentration function, defined by Cifarelli and Regazzini (1987), as a generalisation of the Lorenz curve. Marshall and Olkin (1979, p.5) give the following definition of Lorenz concentration curve (also known as the Lorenz-Gini curve): "Consider a population of n individuals, and let  $x_i$  be the wealth of individual i,  $i = 1, \ldots, n$ . Order the individuals from poorest to richest to obtain  $x_{(1)}, \ldots, x_{(n)}$ . Now plot the

points  $(k/n, S_k/S_n), k = 0, \ldots, n$ , where  $S_0 = 0$  and  $S_k = \sum_{i=1}^n x_{(i)}$  is the total wealth

of the poorest k individuals in the population. Join these points by line segments to obtain a curve connecting the origin with the point  $(1, 1) \cdots$  Notice that if total wealth is uniformly distributed in the population, then the Lorenz curve is a straight line. Otherwise, the curve is convex and lies under the straight line."

The classical definition of concentration refers to the discrepancy between a probability P, which gives mass  $x_{(i)}/S_n$  to  $\theta_i, i=1,\ldots,n$ , and the uniform distribution  $P_0$  on  $\Theta=\{\theta_1,\ldots,\theta_n\}$ . Cifarelli and Regazzini (1987) defined the c.f. of P with respect to (w.r.t.)  $P_0$ , where P and  $P_0$  are two probability measures on the same measurable space  $(\Theta,\mathcal{F})$ . According to the Radon-Nikodym theorem, there is a unique partition  $\{N,N^C\}\subset\mathcal{F}$  of  $\Theta$  and a nonnegative function h on  $N^C$  such that  $P(E)=\int\limits_{E\cap N^C}h(\theta)P_0(d\theta)+P_s(E\cap N),\ \forall\ E\in\mathcal{F},P_0(N)=0,\ P_s(N)=P_s(\Theta),$  where  $P_a(\cdot)=\int\limits_{E\cap N^C}h(\theta)P_0(d\theta)$  and  $P_s$  denote the absolutely continuous and the singular part of P w.r.t.  $P_0$ , respectively. Set  $h(\theta)=\infty$  all over N and define

singular part of P w.r.t.  $P_0$ , respectively. Set  $h(\theta) = \infty$  all over N and define  $H(y) = P_0(\{\theta \in \Theta : h(\theta) \le y\}), c(x) = \inf\{y \in \Re : H(y) \ge x\}$ . Finally, let  $L(x) = \{\theta \in \Theta : h(\theta) \le c(x)\}$  and  $L^-(x) = \{\theta \in \Theta : h(\theta) < c(x)\}$ .

Definition 2. The function  $\varphi:[0,1]\to [0,1]$  is said to be the concentration function of P w.r.t.  $P_0$  if  $\varphi(x)=P(L^-(x))+c(x)\{x-H(c(x)^-)\}$  for  $x\in (0,1), \ \varphi(0)=0$  and  $\varphi(1)=P_a(\Theta)$ .

When the dependence on P is to be emphasized, we will use the notations  $h_P(\theta)$ ,  $H_P(x)$ ,  $c_P(x)$ ,  $L_P(x)$  and  $\varphi_P(x)$ .

Observe that

$$\varphi(x) = \begin{cases} P(L(x)) & x = H(c(x)) = P_0(L(x)) \\ P(L^-(x)) & x = H(c(x)^-) = P_0(L^-(x)) \end{cases}$$

while  $\varphi(x)$  is defined by linear interpolation on  $\{x: H(c(x)) > x < H(c(x))\}$ , if it is not empty. Furthermore, as proved in Cifarelli and Regazzini (1987),  $\varphi(x)$  is a nondecreasing, continuous and convex function such that  $\varphi(x) \equiv 0 \Leftrightarrow P \perp$ 

$$P_0, \ \varphi(x) = x \ \forall \ x \in [0,1] \Leftrightarrow P = P_0 \ \text{and} \ \varphi(x) = \int\limits_0^{c(x)} \{x - H(t)\} dt = \int\limits_0^x c(t) dt.$$

As pointed out by Cifarelli and Regazzini (1987, Remark 2.1), the definition of the c.f. can be extended to bounded positive measures which need not be probability measures. If P is such a measure, its concentration function w.r.t.  $P_0$  coincides with that of  $P_a$  w.r.t.  $P_0$ .

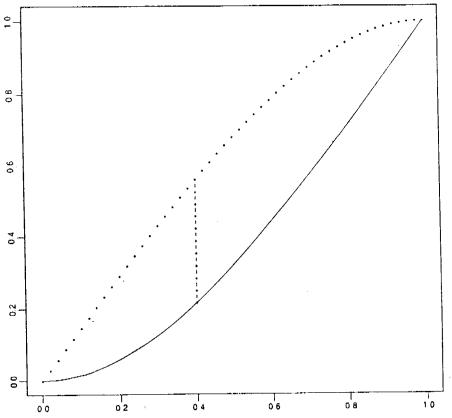


Figure 1. Concentration function  $\varphi(x)(--)$  and  $1-\varphi(1-x)(\cdots)$  of  $P \sim \mathcal{G}(2, 2)$  w.r.t.  $P_0 \sim \mathcal{E}(1)$ .

As an example, the c.f.  $\varphi(x)$  of  $P \sim \mathcal{G}(2,2)$  w.r.t. $P_0 \sim \mathcal{E}(1)$  is plotted in Fig.1, and it is shown, e.g., that [.216, .559] is the range spanned by the probability, under P, of the sets A with  $P_0(A) = .4$ . Such a range is a consequence of the following theorem, due to Cifarelli and Regazzini (1987), which provides an interesting inter-

pretation of the c.f.; in fact, given any  $x \in [0, 1]$ , then the probability, under P, of any A with  $P_0$ -measure x, is such that  $\varphi(x) \leq P(A) \leq 1 - \varphi(1 - x)$ .

Theorem 1. If  $A \in \mathcal{F}$ ,  $P_0(A) = x$ , then  $\varphi(x) \leq P_a(A)$ . Moreover if  $x \in [0,1]$  is adherent to the range of H, then  $B_x$  exists such that  $P_0(B_x) = x$  and

$$\varphi(x) = P_a(B_x) = \min\{P(A) : A \in \mathcal{F} \text{ and } P_0(A) \ge x\}. \tag{2}$$

If  $P_0$  is nonatomic, then (2) holds for any  $x \in [0,1]$ .

Theorem 1 allows to express g-neighbourhoods by means of c.f.'s.

Proposition 1. The set  $K_g = \{P \in \mathcal{P} : \phi_P(x) \geq g(x), \ \forall \ x \in [0,1]\}$  is a g-neighbourhood of  $P_0$  as defined in (1).

#### 3. Topology over $\mathcal{P}$

If G is a suitable class of monotone nondecreasing, convex continuous functions on [0,1], then the class of neighbourhoods  $\{K_g\}_{\{g\in G\}}$  can be used to define a topology on the space  $\mathcal{P}$  of all the probability measures on  $(\Theta, \mathcal{F})$ .

When the dependence on  $P_0$  has to be emphasized,  $\varphi_P^{P_0}$  denotes the c.f. of P w.r.t.  $P_0$  and let  $K_q(P_0) = \{P \in \mathcal{P} : \varphi_P^{P_0}(x) \geq g(x), \forall x \in [0,1]\}.$ 

Proposition 2. Let G be a class of monotone nondecreasing, continuous, convex functions  $g:[0,1] \to [0,1]$ , with g(0)=0 and let G be such that, for any  $g \in G$ , there exists  $\bar{g}, \bar{g} \in G$  such that  $\tilde{g}(g(x)) \geq g(x), \forall \ x \in [0,1]$ . Then there exists a topology T on P such that the class  $\{K_g(P_0)\}_{g \in G}$  is a fundamental system of neighbourhoods of  $P_0$ .

Before proving Proposition 2, we need the following lemma.

Lemma 1. Consider  $\mathcal{N}(P_0) = \{U \subset \mathcal{P} : K_g(P_0) \subseteq U \text{ for some } g \in G\}$ . The following properties hold:

- (1) If  $U_1 \subset U_2$  and  $U_1 \in \mathcal{N}(P_0)$ , then  $U_2 \in \mathcal{N}(P_0)$ .
- (II) If  $U_1, U_2, \ldots, U_n \in \mathcal{N}(P_0)$ , then  $\bigcap_{i=1}^n U_i \in \mathcal{N}(P_0)$ .
- (III) For any  $U \in \mathcal{N}(P_0), P_0 \in U$ .
- (IV) If  $U_1 \in \mathcal{N}(P_0)$ , then there exists  $U_2 \in \mathcal{N}(P_0)$  such that  $U_1 \in \mathcal{N}(P_1)$  for any  $P_1 \in U_2$ .

*Proof.* The proofs of I, II, III are trivial. Given  $U_1 \in \mathcal{N}(P_0)$ , there exists  $g \in G$  such that

$$\{P \in \mathcal{P} : \varphi_P^{P_0}(x) \ge g(x), \ \forall \ x \in [0,1]\} \subseteq U_1.$$

 $\square$ 

To prove IV, it is sufficient to show that there exists  $\tilde{g} \in G$  such that

$$\{P \in \mathcal{P} : \varphi_P^{P_1}(x) \ge \tilde{g}(x), \ \forall \ x \in [0,1]\} \subseteq \{P \in \mathcal{P} : \varphi_P^{P_0}(x) \ge g(x), \ \forall \ x \in [0,1]\} \subseteq U_1, \dots (3)$$

for  $P_1$  belonging to  $K_{\bar{g}}(P_0)$ , and for a suitable  $\bar{g}$ . Take  $\bar{g}$  and  $\bar{g}$  such that  $\bar{g}(\bar{g}(x)) \ge g(x), \forall x \in [0,1]$ , and  $U_2 = K_{\bar{g}}(P_0)$ .

If  $P_1 \in U_2$ , then it follows, from Theorem 1, that

$${A \in \mathcal{F} : P_0(A) \ge x} \subseteq {A \in \mathcal{F} : P_1(A) \ge \bar{g}(x)}.$$

Hence if  $\varphi_P^{P_1}(x) \geq \tilde{g}(x), \ \forall \ x \in [0, 1]$ , then

$$\inf_{\{P_b(A) \ge x\}} P(A) \ge \inf_{\{P_b(A) \ge \tilde{g}(x)\}} P(A) \ge \tilde{g}(g(x)) \ge g(x)$$

which proves (3).

Let us come to the proof of Proposition 2.

*Proof.* From I, II, III, IV it follows that there exists a unique topological structure  $\mathcal{T}$  on P such that, for each  $P_0 \in \mathcal{P}$ ,  $\mathcal{N}(P_0)$  is the set of the neighbourhoods of  $P_0$  in the topology  $\mathcal{T}$ . Moreover the class of neighbourhoods  $\{K_g(P_0)\}_{g\in G}$  is a fundamental system of neighbourhoods of  $P_0$  in  $\mathcal{T}$  (see Bourbaki, 1989).

Example 1. The trivial topology  $T_A$ , in which any probability measure is an open set, is obtained taking any G such that  $g_1 \in G$ , where  $g_1(x) = x, \forall x \in [0,1]$ .

Example 2. Considering  $G = \{g_{\varepsilon}(x) = \max\{0, x - \varepsilon\}, \forall x \in [0, 1], 0 < \varepsilon \leq 1\}$ , the topology  $T_V$  of the variational metric in  $\mathcal{P}$  is obtained. In such a case, all the requirements about the functions  $g_{\varepsilon}$  are satisfied, along with the property  $\tilde{g}(\bar{g}(x)) \geq g_{\varepsilon}(x), \forall x \in [0, 1]$ , for any  $\varepsilon, 0 < \varepsilon \leq 1$ , e.g. taking  $\tilde{g} = g_{\varepsilon_1}, \ \bar{g} = \bar{g}_{\varepsilon_2}$  with  $\varepsilon \geq \varepsilon_1 + \varepsilon_2$ .

Example 3. A topology  $\mathcal{T}_{\alpha}$  is obtained when taking  $G = \{g_{\alpha} : g_{\alpha}(x) = x^{\alpha}, \forall x \in [0,1], 1 < \alpha < \infty\}$ . In such a case, all the requirements about the functions  $g_{\alpha}$  are satisfied, along with the property  $\tilde{g}(g(x)) \geq g_{\alpha}(x), \forall x \in [0,1]$ , for any  $\alpha, 1 < \alpha < \infty$ , e.g. taking  $\tilde{g} = \bar{g} = g_{\sqrt{\alpha}}$ . The topology  $\mathcal{T}_{\alpha}$  is finer than the topology  $\mathcal{T}_{V}$ . In fact, let  $U_{V}$  be a neighbourhood of  $P_{0} \in \mathcal{P}$  in  $\mathcal{T}_{V}$ . It is easy to prove that there exists  $\varepsilon$  such that

$$U_V \supseteq \{P \in \mathcal{P} : \varphi_P^{P_0}(x) \ge \max\{0, x - \varepsilon\}, \ \forall x \in [0, 1]\}.$$

Let  $\alpha > 1$  be such that  $x^{\alpha} \geq \max\{0, x - \epsilon\}$  for any  $x \in [0, 1]$ , then it follows that

$$U_V\supseteq K_{g_{\sigma}}(P_0)$$

so that  $U_V$  is a neighbourhood of  $P_0$  in the topology  $\mathcal{T}_{\alpha}$ . Hence  $\mathcal{T}_{\alpha}$  is finer than  $\mathcal{T}_V$ . It follows that every continuous functional on  $(\mathcal{P}, \mathcal{T}_V)$  is a continuous functional on  $(\mathcal{P}, \mathcal{T}_{\alpha})$ .

As pointed out by Cifarelli and Regazzini (1987), the concentration function w.r.t. a fixed measure  $P_0$  can be used to introduce a partial ordering in the space  $\mathcal{P}$  of all the probability measures on  $(\Theta, \mathcal{F})$ .

Definition 3. The probability measure  $P_1$  is to be said not less concentrated than  $P_2$  w.r.t.  $P_0$  if and only if  $\varphi_{P_1}(x) \leq \varphi_{P_2}(x)$ , for any  $x \in [0,1]$ . We will denote it as  $P_2 \leq P_1$ .

If there exists a probability measure, say  $\bar{P}$ , whose c.f. coincides with g, then the definition of g-neighbourhood can be reformulated as:

$$K_q = \{ P \in \mathcal{P} : P \leq \tilde{P} \}.$$

Such a  $\dot{P}$  exists, provided that g is compatible with  $P_0$ , according to the following definition:

Definition 4. A function  $g:[0,1] \to [0,1]$  is said to be compatible with  $P_0$  if g is a monotone nondecreasing, continuous, convex function, with g(0) = 0, and there exists a correspondence  $\chi: S = \{\theta \in \Theta: P_0(\{\theta\}) \neq 0\} \to [0,1]$  such that, if  $\theta_1 \neq \theta_2$ ,

$$(\chi(\theta_1), \chi(\theta_1) + P_0(\{\theta_1\})) \cap (\chi(\theta_2), \chi(\theta_2) + P_0(\{\theta_2\})) = \emptyset$$

and for any  $\theta \in S$ , it follows that there exist  $c_{\theta}$  and  $d_{\theta}$  such that

$$g_{|(\chi(\theta),\chi(\theta)+P_{\theta}(\{\theta\})}(x)=c_{\theta}x+d_{\theta},$$

where  $g_{|(a,b)}(x)$  is the restriction of g(x) to the interval (a,b).

Observe that if  $P_0$  is nonatomic, then every monotone nondecreasing, continuous, convex function g such that g(0) = 0 is compatible with  $P_0$ .

Lemma 2. Let  $\Theta$  be a Polish space and  $\mathcal{F}$  the  $\sigma$ -algebra of Borel sets. Then there exists a total ordering on  $\Theta$ , denoted by <, such that, for any  $\bar{\theta}$ , the set  $\{\theta \in \Theta : \theta < \bar{\theta}\}$  belongs to  $\mathcal{F}$ .

*Proof.* Let  $\{\theta_n\}_{n\geq 1}$  be a dense subset of  $\Theta$ . Introduce in  $\Theta$  the following ordering:  $\theta \leq \bar{\theta}$  if and only if

$$(d(\theta_1, \theta) < d(\theta_1, \bar{\theta})) \bigvee (d(\theta_n, \theta) = d(\theta_n, \bar{\theta}), n = 1, 2, \ldots)$$

$$\bigvee \left(\bigvee_{n=1}^{\infty} \left(d(\theta_i,\theta) = d(\theta_i,\bar{\theta}), \ i=1,2,\ldots,n, d(\theta_{n+1},\theta) < d(\theta_{n+1},\bar{\theta})\right)\right).$$

If  $\theta \leq \bar{\theta}$  and  $\bar{\theta} \leq \theta$ , then  $d(\theta_n, \theta) = d(\theta_n, \bar{\theta})$ , for any  $n = 1, 2, \ldots$  Since the topology is Hausdorff and  $\{\theta_n\}_{n\geq 1}$  is dense, it follows that  $\theta = \bar{\theta}$ . Moreover the set  $\{\theta \in \Theta : \theta < \bar{\theta}\}$  is a denumerable union of measurable sets and therefore it is measurable.

Theorem 2. Given a function  $g:[0,1] \to [0,1]$ , there exists at least one measure P such that g is the c.f. of P w.r.t.  $P_0$  if and only if g is compatible with  $P_0$ .

Proof. The necessity of the condition is trivial. Let us prove the sufficiency. From the Lemma 2,  $\Theta$  can be endowed with a total ordering, which will be denoted by <, such that, for any  $\tilde{\theta}$ , the set  $\{\theta \in \Theta : \theta < \theta\}$  is measurable. Let  $P_s$  be a measure on  $\Theta$ , singular w.r.t.  $P_0$  and such that  $P_s(\Theta) = 1 - g(1)$ . Let  $\chi$  and S be as in Definition 4. Define  $T = [0,1] \setminus \bigcup_{\theta \in S} (\chi(\theta), \chi(\theta) + P_0(\{\theta\}))$ . The monotone function  $V(x) = x - \sum_{\{\theta \in S; \chi(\theta) < x\}} P_0(\theta)$  determines a one to one correspondence between T

and [0,t], where  $t=P_0(S^C)$ . Hence there exists a measurable monotone function  $W:S^C\to T$  such that

$$P_0(\{\theta \in S^C : \theta < \theta\}) = V(W(\bar{\theta})). \qquad \dots (4)$$

From (4) the measure  $\mu$  defined on every Borel subset A of T as  $\mu(A) = P_0(W^{-1}(A))$  is the restriction of the Lebesgue measure. Since g is a convex function, then it is differentiable almost everywhere. Hence, if  $A_x$  denotes the set where g is not differentiable and  $A = W^{-1}(A_x)$ , then  $P_0(A) = \mu(A_x) = 0$ . Consider the function

$$h(\theta) = \begin{cases} c_{\theta} & \theta \in S \\ g'_{x}(W(\theta)) & \theta \in \Theta \setminus (A \cup S) \\ 0 & \theta \in A. \end{cases}$$

Observe that the restriction of h to  $\Theta \setminus (A \cup S)$  is monotone nondecreasing. Moreover, since W is measurable, then h is measurable. Let  $P_a$  be the measure whose Radon-Nikodym derivative w.r.t.  $P_0$  is h and let  $P = P_s + P_a$ . Then, since h is monotone nondecreasing on  $\Theta \setminus (A \cup S)$  then, if  $x = H(c(x)^-)$ ,

$$\int\limits_{h(\theta)< c_x} h(\theta) P_0(d\theta) = \sum\limits_{\theta \in S; \chi(\theta)< x} c_\theta P_0(\{\theta\}) + \int\limits_{T \cap [0,x]} g'(t) \mu(dt) = g(x).$$

Hence it is easily seen that P is a probability measure and that g is the c.f. of P w.r.t.  $P_0$ .

Observe that, as it results from the proof, the measure, whose concentration function is g, is generally not unique, since it depends on  $P_s$  and the arbitrary ordering on  $\Theta$ . Furthermore, the singular component  $P_s$  is present if and only if g(1) < 1.

#### 4. Representation Theorem

Let  $P_0$  be a nonatomic probability measure on  $\theta$  and let  $g:[0,1] \to [0,1]$  be a function compatible with  $P_0$ . It will be proved now that all the probability measures in a g-neighbourhood  $K_g$  of  $P_0$  are mixtures of the extremal ones in  $E_g$ , where  $E_g = \{P \in \mathcal{P} : \phi_P(x) = g(x)\}$ . Different proofs lead to such a result when either

g(1) = 1 or g(1) < 1; the former case admits just probability measures absolutely continuous w.r.t.  $P_0$  while singularities are admitted in the latter one.

Consider the space  $\mathcal{P}$  of all probability measures on  $\Theta$  endowed with the weak topology; then  $\mathcal{P}$  can be metrized as a complete separable metric space (see Parthasarathy, 1967, pp. 43-46).

Lemma 3. If g(1) = 1, then  $K_g$  is a convex, compact subset of P in the weak topology.

*Proof.*  $K_g$  is convex. Let  $P_1, P_2 \in K_g, 0 < \alpha < 1$  and  $P = \alpha P_1 + (1 - \alpha)P_2$ . If  $A \in \mathcal{F}$  is such that  $P_0(A) \geq x$ , then it follows, from the definition of  $K_g$  and Theorem 1, that

$$P(A) \ge \inf_{P_0(B) \ge x} P(B) \ge \alpha \inf_{P_0(B) \ge x} P_1(B) + (1 - \alpha) \inf_{P_0(B) \ge x} P_2(B) \ge g(x).$$

 $K_g$  is closed. Let  $\{P_n\}_{n\geq 1}$  be a sequence of probability measures in  $K_g$  converging weakly to a probability measure P. For closed sets C we have  $P(C)\geq \lim_{n\to\infty}P_n(C)$ . Hence for any  $\varepsilon>0$  there exists  $n_0(C)$  such that, for any  $n(C)>n_0(C), P_{n(C)}(C)\leq P(C)+\varepsilon$ . Furthermore, let  $P_0(C)=x_0-\delta$ , then it follows that  $g(x_0-\delta)-\varepsilon\leq P(C)$ . Let  $S\in\mathcal{F}$  be such that  $P_0(S)=x_0$ . Then it follows that

$$P(S) = \sup_{\{C \text{ closed } : C \subset S\}} P(C) \ge \sup_{\delta > 0} g(x_0 - \delta) - \epsilon \ge g(x_0) - \epsilon.$$

Hence  $\varphi_P(x) \geq g(x)$ .

 $K_g$  is compact. Since  $K_g$  is closed, it is sufficient to show that for any  $\varepsilon > 0$  there exists a compact set  $R \in \mathcal{F}$  such that for every  $P \in K_g, P(R) \ge 1 - \varepsilon$  (see Parthasarathy, 1967, p. 47). Let  $x_{\varepsilon}$  be such that  $g(x_{\varepsilon}) = 1 - \varepsilon$ . Since  $\Theta$  is a Polish space, it follows (see Parthasarathy, 1967, p. 29) that  $P_0$  is tight, i.e. there exists a compact set  $R \in \mathcal{F}$  such that  $P_0(R) > x_{\varepsilon}$ . Then, applying Theorem 1, it follows that  $P(R) > 1 - \varepsilon$ , for any  $P \in K_g$ .

Let  $P_0$  be a nonatomic probability measure. Consider the set of extremal points of  $K_g$ , that is the probability measures  $P \in K_g$  such that

$$P = \alpha P_1 + (1 - \alpha)P_2, P_1 \in K_g, P_2 \in K_g, 0 < \alpha < 1 \iff P = P_1 = P_2.$$

Proposition 3. The set of all the extremal points of  $K_g$  is contained in  $E_g$ . If g(1) = 1, then it coincides with  $E_g$ .

*Proof.* Suppose g(1) = 1 and let  $P \in E_g$ . If  $P = \alpha P_1 + (1 - \alpha)P_2, P_1, P_2 \in K_g$ , then  $P_1$  and  $P_2$  belong to  $E_g$ . In fact, suppose that  $P_1 \notin E_g$ , so that there exists

 $x \in [0,1]$  such that  $\inf_{P_0(A) \ge x} P_1(A) > g(x)$  and then

$$\inf_{P_0(A)\geq x}P(A)\geq \alpha\inf_{P_0(A)\geq x}P_1(A)+(1-\alpha)\inf_{P_0(A)\geq x}P_2(A)>g(x),$$

because  $P_2 \in K_g$  implies that  $\inf_{P_0(A) \ge x} P_2(A) \ge g(x)$ . Therefore, it follows that  $\varphi_{P_0}(x) = \varphi_{P_0}(x) = g(x)$  and  $c_{P_0}(x) = c_{P_0}(x) = c_{P_0}(x)$  almost everywhere.

Consider  $L_P(x)$  for any  $x \in [0,1]$  such that  $x = H_P(c_P(x))$ . Since  $P, P_1, P_2 \in E_g$ , then  $P(L_P(x)) = \alpha P_1(L_P(x)) + (1-\alpha)P_2(L_P(x))$  implies that for every  $x \in [0,1]$ 

$$P_1(L_P(x)) = P_2(L_P(x)) = P(L_P(x)) = g(x).$$
 (5)

Analogously, we have  $P_1(L_P^-(x)) = P_2(L_P^-(x)) = P(L_P(x))$  if  $x = H_P(c_P(x)^-)$ . If  $x = H_P(c_P(x))$ , then, for any  $\delta > 0$ ,  $c_P(x + \delta) > c_P(x)$  which implies  $c_{P_1}(x + \delta) > c_{P_1}(x)$  and  $c_{P_2}(x + \delta) > c_{P_2}(x)$ . Hence  $x = H_{P_1}(c_{P_1}(x)) = H_{P_2}(c_{P_2}(x))$  so that  $P_1(L_{P_1}(x)) = P_2(L_{P_2}(x)) = g(x)$ . From this and (5) it follows that, for any  $y \ge 0$ ,

$$\{\theta \in \Theta : h_P(\theta) \le y\} = \{\theta \in \Theta : h_{P_1}(\theta) \le y\} = \{\theta \in \Theta : h_{P_2}(\theta) \le y\} \text{ a.s. } -P_0.$$

Analogously,

$$\{ heta \in \Theta: h_P( heta) < y\} = \{ heta \in \Theta: h_{P_1}( heta) < y\} = \{ heta \in \Theta: h_{P_2}( heta) < y\} ext{ a.s. } -P_0.$$

Hence  $P_1 = P_2 = P$  and every probability measure in  $E_g$  is an extremal point for  $K_g$ .

Let now P be a probability measure not belonging to  $E_g$ . Then there exists  $x \in [0,1]$  such that  $\varphi_P(x) > g(x)$ . Since  $\varphi_p$  and g are continuous, there exists a neighbourhood  $\mathcal{U}$  of x such that  $\varphi_P(x) > g(x)$  for every x in  $\mathcal{U}$ .

Let  $(x_1, x_2)$  be the largest interval, eventually (0, 1), such that  $x \in (x_1, x_2)$  implies  $\varphi_P(x) > g(x)$ . Since  $\varphi_P$  and g are continuous,  $\varphi_P(x_1) = g(x_1)$  and  $\varphi_P(x_2) = g(x_2)$ .

Let  $\bar{x} \in (x_1, x_2)$  and let  $a = (\varphi_P)'_+(\bar{x})$ . Then  $g'_-(x_1) < a < g'_+(x_2)$ . It is easily seen that  $a_1, a_2, c_{a_1}, c_{a_2}$  can be chosen such that  $g'_-(x_1) < a_1 < a < a_2 < g'_+(x_2)$  and

$$\bar{\varphi}(x) = \begin{cases} \varphi_P(x) & (\varphi_P)'_+(x) \le a_1 \text{ or } (\varphi_P)'_+(x_2) > a_2 \\ a_1x + c_{a_1} & a_1 < (\varphi_P)'_-(x) \le (\varphi_P)'_+(x) \le a \\ a_2x + c_{a_2} & a < (\varphi_P)'_-(x) \le (\varphi_P)'_+(x) \le a_2 \end{cases}$$

is a continuous function with the property  $g(x) \leq \bar{\varphi}(x) \leq \varphi_P(x)$  for any  $x \in [0,1]$ . Let  $\epsilon = \min(a_2 - a, a - a_1)$ . Since  $P_0$  is nonatomic, the set  $\{\theta \in \Theta : h_P(\theta) \in [a - \epsilon/3, a + \epsilon/3]\}$  contains more than one point. Hence there exists a non-constant function  $\delta$  defined on  $\Theta$  such that  $\delta(\theta) = 0$  if  $h_P(\theta) \notin [a - \epsilon/3, a + \epsilon/3], |\delta(\theta)| < \epsilon/3$  and  $\int_{\Theta} \delta(\theta) P_0(d\theta) = 0$ . Define  $h_1(\theta) = h_P(\theta) + \delta(\theta), h_2(\theta) = h_P(\theta) - \delta(\theta)$ . Hence  $h_i(\theta) = h_P(\theta)$  if  $h_P(\theta) \notin [a_1, a_2], a - \epsilon < h_i(\theta) < a + \epsilon$  if  $a - \epsilon < h_P(\theta) < a + \epsilon$ 

and  $\int_{\{h_i(\theta) \leq a - \epsilon\}} h_i(\theta) P_0(d\theta) = \int_{\{h_P(\theta) \leq a - \epsilon\}} h_P(\theta) P_0(d\theta)$ , i = 1, 2. Define  $P_i$  as the probability measure on  $\Theta$  whose Radon-Nikodym derivative with respect to  $P_0$  is  $h_i$ , i = 1, 2. It follows that  $\varphi_{P_1}$  and  $\varphi_{P_2}$  satisfy the condition

$$\varphi(x) = \varphi_P(x) \text{ for } (\varphi_P)'_+(x) < a_1 \text{ and } (\varphi_P)'_+(x) > a_2. \qquad \ldots (6)$$

Moreover, since  $\bar{\phi}$  is the most concentrated curve among those satisfying (6),

$$\varphi_{P_{\epsilon}}(x) \ge \bar{\varphi}(x) > g(x). \qquad \dots (7)$$

Since  $P = (P_1 + P_2)/2$  and (7) holds, then P is not an extremal point of  $K_g$ .

Suppose now g(1) < 1; it can be proved, as before, that  $E_g$  is an extremal subset, i.e. that  $P \in E_g$  is a convex combination just of  $P_1, P_2 \in E_g$ . No probability measure  $P \notin E_g$  is an extremal one, because it can be expressed as a combination of two other measures, as before, whose c.f.'s are above g. The extremal points, if any, are thus in  $E_g$ .

Every probability measure whose c.f. is greater than g can be represented as a mixture of probability measures having g as c.f.

Theorem 3. Let the function  $g:[0,1] \to [0,1]$  be compatible with a nonatomic probability measure  $P_0$ . For any probability measure  $P \in K_g$ , there exists a probability measure  $\mu_{\bar{P}}$  on  $\mathcal{P}$  such that  $\mu_{\bar{P}}(E_g) = 1$  and  $\bar{P} = \int_{\mathcal{P}} P \mu_{\bar{P}}(dP)$ .

**Proof.** Suppose g(1) = 1 and let  $K_g$  be the set of the probability measures P such that  $\varphi_P(x) \geq g(x)$ . It was proved in Lemma 3 that  $K_g$  is convex and compact in P. Moreover in Proposition 3 it was proved that  $E_g$  is the extremal set of  $K_g$ .

Consider the topological vector space  $C(\Theta)$  of all bounded continuous functions on  $\Theta$ , endowed with the supremum topology, and let C' be its dual space; it can be easily seen that  $\mathcal{P} \subset C'$ .

The  $C(\Theta)$ -topology of C' is a locally convex vector topology on C' (see Rudin, 1991, p. 68). Since  $K_g$  is metrizable, because  $\mathcal{P}$  is, and it is also convex and compact in C', then the Choquet's theorem (Phelps, 1966, p. 19) implies the result.

Suppose now  $g(1)=1-\varepsilon<1$ . Let  $\bar{P}_a(\Theta)=1-\eta$ , where, for any  $P,\bar{P}_a$  denote the absolutely continuous part of P with respect to  $P_0$ . If  $\eta=\varepsilon$ , then the previous proof is applied to the set  $\mathcal{P}_{1-\varepsilon}$  of the measures P on  $\Theta$  such that  $P(\Theta)=1-\varepsilon$ , proving that there exists a probability measure  $\mu$  on  $\mathcal{P}_{1-\varepsilon}$  such that  $\bar{P}_a=\int\limits_{\{P\in\mathcal{P}_{1-\varepsilon},\varphi_P=g\}}P_{\mu}(dP)$ . Hence

$$\bar{P} = \bar{P}_s + \int\limits_{\{P \in \mathcal{P}_{1-\epsilon}, \varphi_P = g\}} P_{\mu}(dP) = \int\limits_{\{P \in \mathcal{P}, \varphi_P = g\}} P_{\bar{\mu}}(dP).$$

Consider the case  $\eta < \varepsilon$ . Let  $\delta > 0$  be such that  $\varphi_{\tilde{P}}(x) > g(x)$  for every  $x \in (1-\delta, 1]$ . Let  $\tilde{g}: [0, 1] \to [0, 1-\eta]$  be a continuous convex monotone nondecreasing function

such that  $\tilde{g}(1) = 1 - \eta$  and

$$\left\{ \begin{array}{ll} \tilde{g}(x) = g(x) & 0 \leq x \leq 1 - \delta \\ g(x) \leq \tilde{g}(x) \leq \varphi_P(x) & 1 - \delta \leq x \leq 1 \end{array} \right.$$

Let  $\tilde{K}_g = \{P: P(\Theta) = 1 - \eta, \ \tilde{g}(x) \leq \varphi_P(x) \leq (1 - \eta)x\}$ . Then it can be proved as in Lemma 3 that  $\tilde{K}_g$  is a convex compact subset of the space  $\mathcal{P}_{(1-\eta)}$  of the measures P on  $\Theta$  such that  $P(\Theta) = 1 - \eta$  endowed with the weak topology. Moreover  $P \in \tilde{K}_g$  is an extremal point if and only if  $\varphi_P(x) = \tilde{g}(x)$  (see Proposition 3). Since  $\bar{P}_a \in \tilde{K}_g$ , there exists a probability measure  $\mu$  on  $\mathcal{P}_{1-\eta}$  such that

$$ar{P}_a = \int\limits_{\{P(\Theta)=1-\eta,\;arphi_P(x)=ar{q}(x)\}} P_\mu(dP).$$

Let P be such that  $P(\Theta)=1-\eta$  and  $\varphi_P(x)=\tilde{g}$  and let h be its Radon-Nikodym derivative with respect to  $P_0$ . It is easy to see that there exist  $\delta$  and  $h_g$  such that  $h_g(\theta)=h(\theta)$  for  $\theta\in L_P(1-\delta), h_g(\theta)< h(\theta)$  for  $\theta\in\Theta\backslash L_P(1-\delta), P_g(\Theta)=1-\varepsilon$  and  $\varphi_{P_g}(x)=g(x)$  for any  $x\in[0,1]$ , where  $h_g$  is the Radon-Nikodym derivative of  $P_g$  with respect to  $P_0$ .

Since  $P-P_g$  is a positive measure on  $\Theta$ , it is well-known that there exists a probability measure Q on  $\Theta$  such that  $P-P_g=(\varepsilon-\eta)\int_{\Theta}\delta_{\theta}Q(d\theta)$  where  $\delta_{\theta}$  is the Dirac measure on  $\theta$ . Hence  $P=P_g+\int_{\Theta}(\varepsilon-\eta)\delta_{\theta}Q(d\theta)$  and

$$\begin{split} \bar{P}_a &= \int\limits_{\{(P_g,Q)\}} \int\limits_{\Theta} [P_g + (\varepsilon - \eta)\delta_\theta] Q(d\theta) \mu'(dP_g,dQ) \\ &= \int\limits_{\{P(\theta):=1-\eta, \, \varphi_g=g\}} P_{\mu_P}(dP). \end{split}$$

Hence 
$$\bar{P} = \int_{\{P(\Theta)=1, \varphi_{\Gamma}=g\}} P_{\mu_{\Gamma}}(DP)$$
.

Observe that if  $P_0$  is nonatomic, then Thoerem 3 does not hold in general. Example 4. Suppose  $\Theta = \{\theta_1, \theta_2, \theta_3\}$ ; let  $P_0(\theta_1) = 0$ ,  $P_0(\theta_2) = P_0(\theta_3) = 1/2$ ,  $P(\theta_1) = 0$ ,  $P(\theta_2) = 1/3$ ,  $P(\theta_3) = 2/3$  and g(x) = 2x/3. Let  $\mu$  be a probability measure on P such that  $\mu(E_q) = 1$  and let

$$P_{\mu} = \int_{\mathcal{D}} P_{\mu}(dP).$$

Since  $P(\theta_1) = 1/3$  for any  $P \in E_g$ , then  $P_1$  can not be represented as a mixture of measures belonging to  $E_g$ .

## 5. Bounds on functionals

Theorem 3 can be applied to show that the supremum (or the infimum) over  $K_g$  of a large class of functionals is equal to the supremum (or the infimum) over  $E_g$ . The proof can be obtained by slightly modifying the one in Sivaganesan and Berger (1989).

Theorem 4. Let f and g be real-valued functions on  $\Theta$  such that  $\int\limits_{\Theta}|f(\theta)|P(d\theta)<\infty$  and  $0<\int\limits_{\Theta}g(\theta)P(d\theta)<\infty$  for any  $P\in K_g$ . Then

$$\sup_{P \in K_{\theta}} \frac{\int_{\Theta} f(\theta) P(d\theta)}{\int_{\Theta} g(\theta) P(d\theta)} = \sup_{P \in E_{\theta}} \frac{\int_{\Theta} f(\theta) P(d\theta)}{\int_{\Theta} g(\theta) P(d\theta)}. \tag{8}$$

The same result holds with "sup" replaced by "inf".

Observe that if  $f(\theta) = A + f_1(\theta)$  and  $g(\theta) = B + g_1(\theta)$ , then (8) becomes

$$\sup_{P \in \mathcal{K}_q} \frac{A + \int_{\Theta} f_1(\theta) P(d\theta)}{B + \int_{\Theta} g_1(\theta) P(d\theta)} = \sup_{P \in \mathcal{E}_q} \frac{A + \int_{\Theta} f(\theta) P(d\theta)}{B + \int_{\Theta} g(\theta) P(d\theta)} \quad . \tag{9}$$

If  $g(\theta) \equiv c$  for some c > 0, then the computation of

$$\sup_{P \in \mathcal{K}_q} \frac{\int_{\Theta} f(\theta) P(d\theta)}{\int_{\Theta} g(\theta) P(d\theta)} = \frac{1}{c} \sup_{P \in \mathcal{E}_q} \int_{\Theta} f(\theta) P(d\theta)$$

becomes easy.

Theorem 5. Let  $H_f(y) = P_0(\{\theta \in \Theta : f(\theta) \le y\}), c_f(x) = \inf\{y : H_f(y) \ge x\}.$   $Then \sup_{P \in K_q} \int_{\Theta} f(\theta)P(d\theta) = \int_0^1 c_f(x)c(x)dx, where c(x) = g'(x) \text{ a.e.}$ 

Proof. Observe that  $\sup_{P \in K_g} \int_{\Theta} f(\theta) P(d\theta) = \sup_{P \in K_g} \int_{\Theta} f(\theta) h_P(\theta) P_0(d\theta)$ . Since  $c_f(x)$  and c(x) are rearrangements (see Hardy-Littlewood-Pólya, 1988, pp. 276-278) of  $f(\theta)$  and  $h(\theta)$ , then  $\int_{\Theta} f(\theta) P(d\theta) \leq \int_{0}^{1} c_f(x) c(x) dx$ . On the other hand there exists  $P \in E_g$  such that  $f(\theta_1) \leq f(\theta_2)$  implies  $h_P(\theta_1) \leq h_P(\theta_2)$ . For such a  $P, \int_{\Theta} f(\theta) P(d\theta) = \int_{0}^{1} c_f(x) c(x) dx$ , which proves the Theorem.

Since  $\inf_{P \in K_g} \int_{\Theta} -f(\theta)P(d\theta) = -\sup_{P \in K_g} \int_{\Theta} f(\theta)P(d\theta)$ , the following corollary to Theorem 5 holds.

Corollary 1. 
$$\inf_{P \in K_g} \int_{\Theta} f(\theta) P(d\theta) = -\int_{0}^{1} c_{-f}(x) c(x) dx.$$

#### 6. Discussion

In this paper we have defined neighbourhoods of a probability measure  $P_0$  and then, by means of a representation theorem, we have been able to provide results on the bounds of functionals defined in such neighbourhoods. The results presented can be used both in economics and in statistics, and not only in the problems in which the concentration curve has been widely used.

We just mention that the Lorenz curve, as well as the related Gini's index, has been sometimes used as a tool to examine how far actual situations are from ideal ones (e.g. to check the fairness in allocating seats in a U.S. legislature, so that representatives are elected by equivalent numbers of voters). In such problems, the function g could be considered as the maximum allowed distance from a uniform probability measure, expressed by means of c.f.'s, obviously.

Fortini and Ruggeri (1994) already applied the results in this paper to the robust Bayesian analysis, where the authors faced the problems of building a class of prior measures in a neighbourhood of a given one and checking if inferences lead to either posterior measures close to a base one or posterior functionals quite insensitive to the changes in the prior.

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